



**Stochastic and non-ODE epidemiological  
models**

**Potchefstroom – Course 03**

**Julien Arino**

**27 February 2024**

# Outline

Sojourn times in compartments

Discrete-time Markov chains

Continuous time Markov chains

Structuration in age

# Pierre Magal (1968 – 2024-02-20)



Sojourn times in compartments

Discrete-time Markov chains

Continuous time Markov chains

Structuration in age

See in particular the work of Horst Thieme

If one considers time of sojourn in compartments from a more detailed perspective, one obtains integro-differential models

We use here continuous random variables. See chapters 12 and 13 in Thieme's book for arbitrary distributions

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## Time to events

We suppose that a system can be in two states,  $S_0$  and  $S_2$

- ▶ At time  $t = 0$ , the system is in state  $S_0$
- ▶ An event happens at some time  $t = \tau$ , which triggers the switch from state  $S_0$  to state  $S_1$

Let us call  $T$  the random variable

*“time spent in state  $S_0$  before switching into state  $S_1$ ”*

The states can be anything:

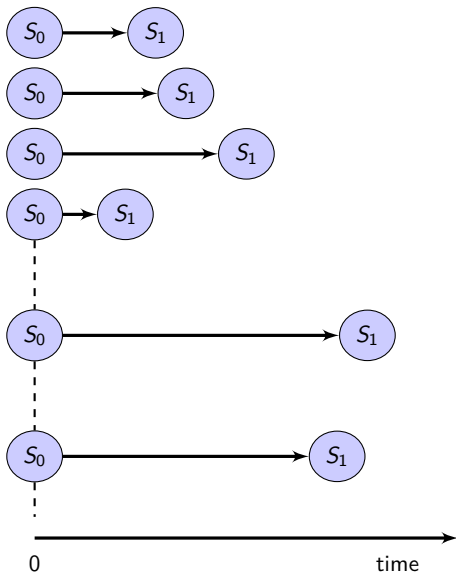
- ▶  $S_0$ : working,  $S_1$ : broken
- ▶  $S_0$ : infected,  $S_1$ : recovered
- ▶  $S_0$ : alive,  $S_1$ : dead
- ▶ ...

We take a collection of objects or individuals that are in state  $S_0$  and want some law for the **distribution** of the times spent in  $S_0$ , i.e., a law for  $T$

For example, we make light bulbs and would like to tell our customers that on average, our light bulbs last 200 years...

We conduct an **infinite** number of experiments, and observe the time that it takes, in every experiment, to switch from  $S_0$  to  $S_1$





# A distribution of probability is a model

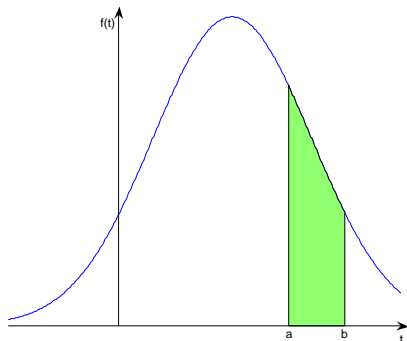
From the sequence of experiments, we deduce a model, which in this context is called a **probability distribution**

We assume that  $T$  is a **continuous** random variable

# Probability density function

Since  $T$  is continuous, it has a continuous **probability density function**  $f$

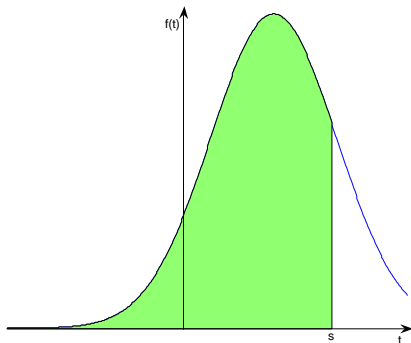
- ▶  $f \geq 0$
- ▶  $\int_{-\infty}^{+\infty} f(s) ds = 1$
- ▶  $\mathbb{P}(a \leq T \leq b) = \int_a^b f(t) dt$



## Cumulative distribution function

The cumulative distribution function (c.d.f.) is a function  $F(t)$  that characterizes the distribution of  $T$ , and defined by

$$F(s) = \mathbb{P}(T \leq s) = \int_{-\infty}^s f(x) dx$$



## Survival function

Another characterization of the distribution of the random variable  $T$  is through the **survival** (or **sojourn**) function

The survival function of state  $S_0$  is given by

$$S(t) = 1 - F(t) = \mathbb{P}(T > t) \quad (1)$$

This gives a description of the **sojourn time** of a system in a particular state (the time spent in the state)

$S$  is a nonincreasing function (since  $S = 1 - F$  with  $F$  a c.d.f.), and  $S(0) = 1$  (since  $T$  is a nonnegative random variable)

The **average sojourn time**  $\tau$  in state  $S_0$  is given by

$$\tau = E(T) = \int_0^{\infty} t f(t) dt$$

Since  $\lim_{t \rightarrow \infty} tS(t) = 0$ , it follows that

$$\tau = \int_0^{\infty} S(t) dt$$

**Expected future lifetime:**

$$\frac{1}{S(t_0)} \int_0^{\infty} t f(t + t_0) dt$$

$$\begin{aligned} S(t) - S(a) &= \mathbb{P} \{ \text{survive during } (a, t) \text{ having survived until } a \} \\ &= \exp \left( - \int_a^t h(u) du \right) \end{aligned}$$

## Hazard rate

The **hazard rate** (or **failure rate**) is

$$\begin{aligned}h(t) &= \lim_{\Delta t \rightarrow 0} \frac{\mathcal{S}(t) - \mathcal{S}(t + \Delta t)}{\Delta t} \\&= \lim_{\Delta t \rightarrow 0} \frac{\mathbb{P}T < t + \Delta t | T \geq t}{\Delta t} \\&= \frac{f(t)}{\mathcal{S}(t)}\end{aligned}$$

It gives probability of failure between  $t$  and  $\Delta t$ , given survival to  $t$ .

We have

$$h(t) = -\frac{d}{dt} \ln \mathcal{S}(t)$$

## Competing risks

Suppose now that the system starts in state  $A$  at time  $t = 0$  and that depending on which of the two events  $\mathcal{E}_1$  or  $\mathcal{E}_2$  takes place first, it switches to state  $B_1$  or  $B_2$ , respectively

Consider the random variables  $T_A$ , *time spent* in state  $A$  (or sojourn time in  $A$ ),  $T_{AB_1}$ , *time before switch to  $B_1$*  and  $T_{AB_2}$ , *time before switch to  $B_2$*

If we consider state  $A$ , we cannot observe the variables  $T_{AB_1}$  or  $T_{AB_2}$ . What is observable is the sojourn time in  $A$

$$T_A^* = \min(T_{AB_1}, T_{AB_2})$$

(where  $*$  indicates that a quantity is observable)



## Failure rate by type of event

We have two (or more) types of events whose individual failure rates have to be accounted for

$$h_j(t) = \lim_{\Delta t \rightarrow 0} \frac{\mathbb{P}(T < t + \Delta t, S = S_j | T \geq t)}{\Delta t}$$

where  $\mathbb{P}(T < t + \Delta t, S = S_j | T \geq t)$  is the probability of failure due to cause  $S_j$  ( $j = 1, 2$  ici), i.e.,  $S$  is a discrete r.v. representing the event that is taking place

By the law of total probability, since only one of the event can take place, if there are  $n$  risks, then

$$h(t) = \sum_{i=1}^n h_j(t)$$

or, identically,

$$S(t) = \exp \left( - \int_0^t \sum_{j=1}^n h_j(s) ds \right)$$

As a consequence, suppose a process is subject to two competing exponential risks with respective distributions with parameters  $\theta_1$  and  $\theta_2$

Then the mean sojourn time in the initial state before being affected by one of the two risks is

$$\frac{1}{\theta_1 + \theta_2}$$

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## The exponential distribution

The random variable  $T$  has an **exponential** distribution if its probability density function takes the form

$$f(t) = \begin{cases} 0 & \text{if } t < 0, \\ \theta e^{-\theta t} & \text{if } t \geq 0, \end{cases} \quad (2)$$

with  $\theta > 0$ . Then the survival function for state  $S_0$  is of the form  $\mathcal{S}(t) = e^{-\theta t}$ , for  $t \geq 0$ , and the average sojourn time in state  $S_0$  is

$$\tau = \int_0^{\infty} e^{-\theta t} dt = \frac{1}{\theta}$$

## Particularities of the exponential distribution

The standard deviation of an exponential distribution is also  $1/\theta$ .  
When estimating  $\theta$ , it is impossible to distinguish the mean and the standard deviation

The exponential distribution is **memoryless**: its conditional probability obeys

$$P(T > s + t \mid T > s) = P(T > t), \quad \forall s, t \geq 0$$

The exponential and geometric distributions are the only memoryless probability distributions

The exponential distribution has a constant hazard function

## The Dirac delta distribution

If for some constant  $\omega > 0$ ,

$$S(t) = \begin{cases} 1, & 0 \leq t \leq \omega \\ 0, & \omega < t \end{cases}$$

meaning that  $T$  has a Dirac delta distribution  $\delta_\omega(t)$ , then the average sojourn time is

$$\tau = \int_0^\omega dt = \omega$$

with standard deviation  $\sigma = 0$

# The Gamma distribution

R.v.  $X$  is **Gamma** distributed ( $X \sim \Gamma(k, \theta)$ ) with **shape parameter**  $k$  and **scale parameter**  $\theta$  (or **rate**  $\beta = 1/\theta$ ) (all positive) if its probability density function takes the form

$$f(x; k, \theta) = \frac{x^{k-1} e^{-\frac{x}{\theta}}}{\Gamma(k)\theta^k} \quad (3)$$

where  $x > 0$  and  $\Gamma$  is the Euler Gamma function, defined for all  $z \in \mathbb{C}$  s.t.  $\text{Re}(z) > 0$  by

$$\Gamma : z \mapsto \int_0^{+\infty} t^{z-1} e^{-t} dt$$



## Properties of the Gamma distribution

Mean  $k\theta$ , variance  $k\theta^2$

Survival function

$$S(t) = 1 - \frac{1}{\Gamma(k)} \gamma\left(k, \frac{t}{\theta}\right) = 1 - \frac{1}{\Gamma(k)} \gamma(k, \beta t)$$

where

$$\gamma(a, x) = \int_0^x t^{a-1} e^{-t} dt$$

is an incomplete Gamma function

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## A model for a cohort with one cause of death

Consider a **cohort** of individuals born at the same time, e.g., the same year

- ▶ At time  $t = 0$ , there are initially  $N_0 > 0$  individuals
- ▶ All causes of death are compounded together
- ▶ The time until death, for a given individual, is a random variable  $T$ , with continuous probability density distribution  $f(t)$  and survival function  $P(t)$

$N(t)$  the cohort population at time  $t \geq 0$

$$N(t) = N_0 P(t) \quad (4)$$

$N_0 P(t)$  proportion of initial population still alive at time  $t$

## Case where $T$ is exponentially distributed

Suppose that  $T$  has an exponential distribution with mean  $1/d$  (or parameter  $d$ ),  $f(t) = de^{-dt}$ . Then the survival function is  $P(t) = e^{-dt}$ , and (4) takes the form

$$N(t) = N_0 e^{-dt} \quad (5)$$

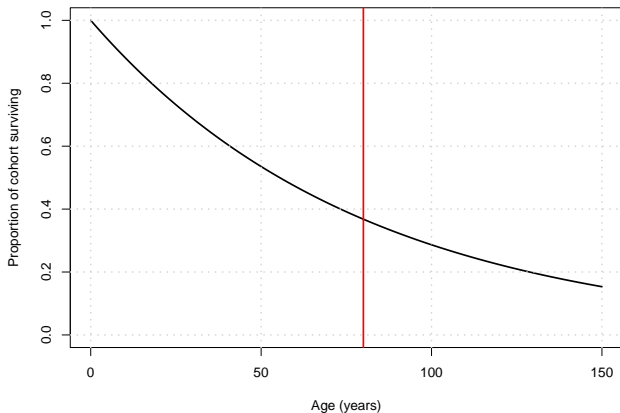
Now note that

$$\begin{aligned} \frac{d}{dt} N(t) &= -dN_0 e^{-dt} \\ &= -dN(t) \end{aligned}$$

with  $N(0) = N_0$ .

⇒ The ODE  $N' = -dN$  makes the assumption that the life expectancy at birth is exponentially distributed

Survival function,  $\mathcal{S}(t) = \mathbb{P}(T > t)$ , for an exponential distribution with mean 80 years



## Case where $T$ has a Dirac delta distribution

Suppose that  $T$  has a Dirac delta distribution at  $t = \omega$ , giving the survival function

$$P(t) = \begin{cases} 1, & 0 \leq t \leq \omega \\ 0, & t > \omega \end{cases}$$

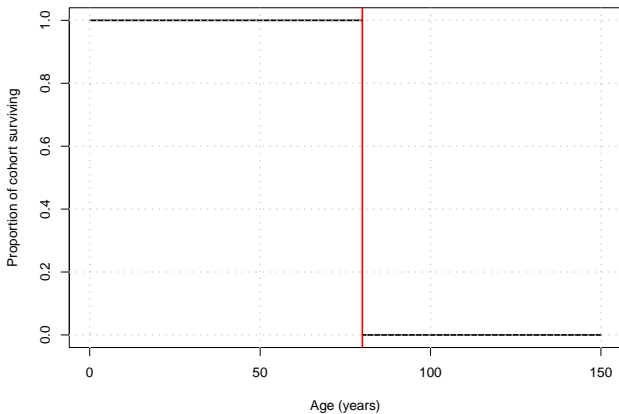
Then (4) takes the form

$$N(t) = \begin{cases} N_0, & 0 \leq t \leq \omega \\ 0, & t > \omega \end{cases} \quad (6)$$

All individuals survive until time  $\omega$ , then they all die at time  $\omega$

Here,  $N' = 0$  everywhere except at  $t = \omega$ , where it is undefined

Survival function,  $\mathcal{S}(t) = \mathbb{P}(T > t)$ , for a Dirac distribution with mean 80 years



## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

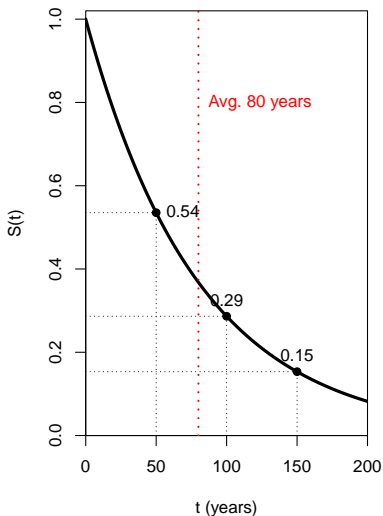
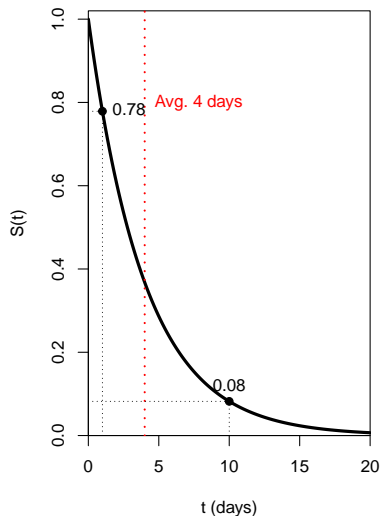
A model with vaccination

Age of vaccination

Conclusions on sojourn times



# Survival for the exponential distribution



## Issues with the exponential distribution

- ▶ Survival drops quickly
- ▶ Survival continues way beyond the mean

Acceptable if what matters is the average duration of sojourn in a compartment (e.g., long term dynamics)

More iffy if one is interested in short-term dynamics

- ▶ Exponential distribution with parameter  $\theta$  has same mean and standard deviation  $1/\theta$ , i.e., a single parameter controls mean and dispersion about the mean

## Exponential distributions are bad but also cool

$X_1$  and  $X_2$  2 i.i.d. (independent and identically distributed) r.v. with parameters  $\theta_1$  and  $\theta_2$ . Then the probability density function of the r.v.  $Z = X_1 + X_2$  is given by the convolution

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^{\infty} f_{X_1}(x_1) f_{X_2}(z - x_1) dx_1 \\ &= \int_0^z \theta_1 e^{-\theta_1 x_1} \theta_2 e^{-\theta_2(z-x_1)} dx_1 \\ &= \theta_1 \theta_2 e^{-\theta_2 z} \int_0^z e^{(\theta_2 - \theta_1)x_1} dx_1 \\ &= \begin{cases} \frac{\theta_1 \theta_2}{\theta_2 - \theta_1} (e^{-\theta_1 z} - e^{-\theta_2 z}) & \text{if } \theta_1 \neq \theta_2 \\ \theta^2 z e^{-\theta z} & \text{if } \theta_1 = \theta_2 =: \theta \end{cases} \end{aligned} \quad (7)$$

## The tool we use

### Theorem 1

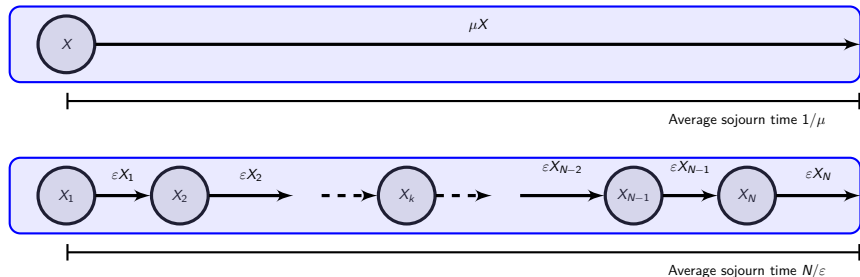
*Let  $X_i$  be independent exponentially distributed random variables with parameter  $\xi$  and  $Y = \sum_{i=1}^n X_i$*

*Then the random variable  $Y \rightsquigarrow E(n, \xi)$ , an Erlang distribution with shape parameter  $n$  and scale parameter  $\xi$*

(Erlang distribution: Gamma distribution with integer shape parameter)

## Consequences for compartmental models

If  $n$  compartments are traversed successively by individuals, with each compartment having an outflow rate of  $1/\xi$  (or a mean sojourn time of  $\xi$ ), then the time of sojourn from entry into the first compartment to exit from the last is Erlang distributed with mean  $E(Y) = n\xi$  and variance  $\text{Var}(Y) = n\xi^2$



I have a Shiny app for this :)

## Example: EVD incubation periods

Consider the incubation period for Ebola Virus Disease. During the 2014 EVD crisis in Western Africa, the WHO Ebola Response Team estimated incubation periods in a 2015 paper

Table S2 in the Supplementary Information in that paper gives the best fit for the distribution of incubation periods for EVD as a Gamma distribution with mean 10.3 days and standard deviation 8.2, i.e.,  $n\varepsilon = 10.3$  and  $\varepsilon\sqrt{n} = 8.2$

From this,  $\varepsilon = 8.2^2/10.3 \simeq 6.53$  and  $n = 10.3^2/8.2^2 \simeq 1.57$ . However, that is a Gamma distribution

## Switching to a compartmental model approach

To use multiple compartments to better fit residence times, we need to find the closest possible Erlang distribution to this Gamma distribution

⇒ compute RSS errors between data points generated from the given Gamma distribution and an Erlang

```
> error_Gamma <- function(theta,shape,t,d) {  
+   test_points <- dgamma(t, shape = shape, scale = theta)  
+   ls_error <- sum((d-test_points)^2)  
+   return(ls_error)  
+ }
```

```

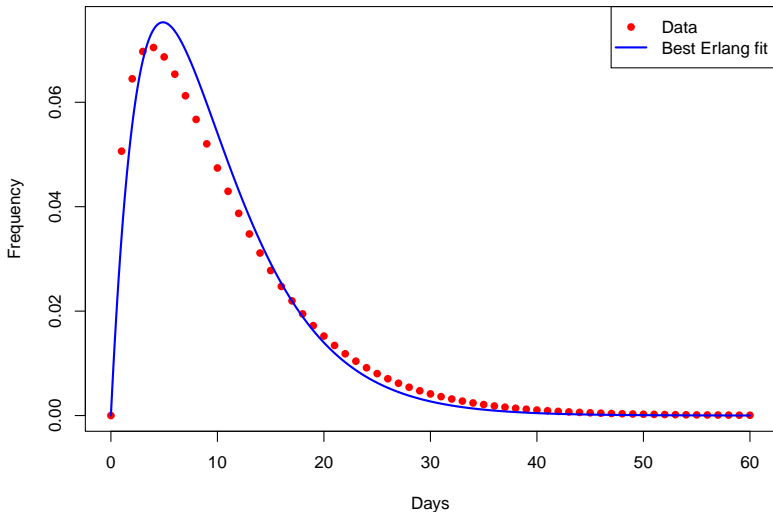
> optimize_gamma <- function(t,d) {
+   max_shape <- 10
+   error_vector <- mat.or.vec(max_shape,1)
+   scale_vector <- mat.or.vec(max_shape,1)
+   for (i in 1:max_shape) {
+     result_optim <- try(optim(par = 3,
+                               fn = error_Gamma,
+                               lower = 0,
+                               method = "L-BFGS-B",
+                               shape = i,
+                               t = t,
+                               d = d),
+                          TRUE)
+     if (!inherits(result_optim,"try-error")) {
+       error_vector[i] <- result_optim$value
+       scale_vector[i] <- result_optim$par
+     } else {
+       error_vector[i] <- NaN
+       scale_vector[i] <- NaN

```



```
> time_points <- seq(0,60)
> data_points <- dgamma(time_points, shape = 1.57,
+                        scale = 6.53)
> # Run the minimization
> optim_fits <- optimize_gamma(time_points,data_points)
> # Which is the best Erlang to fit the data
> idx_best <- which.min(optim_fits$error)
```

Now plot the result as well as the original curve (code chunk not shown)



## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

# An SIS model

## Hypotheses

- ▶ Individuals typically recover from the disease
- ▶ The disease does not confer immunity
- ▶ There is no birth or death (from the disease or natural)  
⇒ Constant total population  $N \equiv N(t) = S(t) + I(t)$
- ▶ Infection is of **standard incidence** type

# Recovery

- ▶ Traditional models suppose that recovery occurs with rate constant  $\gamma$
- ▶ Here, of the individuals that become infective at time  $t_0$ , a fraction  $P(t - t_0)$  remain infective at time  $t \geq t_0$
- ▶  $\Rightarrow$  For  $t \geq 0$ ,  $P(t)$  is a survival function. As such, it verifies  $P(0) = 1$  and  $P$  is nonnegative and nonincreasing

## Model for infectious individuals

Since  $N$  is constant,  $S(t) = N - I(t)$  and we need only consider the following equation (where  $S$  is used for clarity)

$$I(t) = I_0(t) + \int_0^t \beta \frac{S(u)I(u)}{N} P(t-u) du \quad (8)$$

- ▶  $I_0(t)$  number of individuals who were infective at time  $t = 0$  and still are at time  $t$ 
  - ▶  $I_0(t)$  is nonnegative, nonincreasing, and such that  $\lim_{t \rightarrow \infty} I_0(t) = 0$
- ▶  $P(t-u)$  proportion of individuals who became infective at time  $u$  and who still are at time  $t$

## Expression under the integral

Integral equation for the number of infective individuals:

$$I(t) = I_0(t) + \int_0^t \beta \frac{(N - I(u))I(u)}{N} P(t - u) du \quad (8)$$

The term

$$\beta \frac{(N - I(u))I(u)}{N} P(t - u)$$

- ▶  $\beta(N - I(u))I(u)/N$  is the rate at which new infectives are created, at time  $u$
- ▶ multiplying by  $P(t - u)$  gives the proportion of those who became infectives at time  $u$  and who still are at time  $t$

Summing over  $[0, t]$  gives the number of infective individuals at time  $t$

## Case of an exponentially distributed time to recovery

Suppose  $P(t)$  such that sojourn time in the infective state has exponential distribution with mean  $1/\gamma$ , i.e.,  $P(t) = e^{-\gamma t}$

Initial condition function  $I_0(t)$  takes the form

$$I_0(t) = I_0(0)e^{-\gamma t}$$

with  $I_0(0)$  the number of infective individuals at time  $t = 0$ .  
Obtained by considering the cohort of initially infectious individuals, giving a model such as (4)

Equation (8) becomes

$$I(t) = I_0(0)e^{-\gamma t} + \int_0^t \beta \frac{(N - I(u))I(u)}{N} e^{-\gamma(t-u)} du \quad (9)$$



Taking the time derivative of (9) yields

$$\begin{aligned} I'(t) &= -\gamma I_0(0)e^{-\gamma t} - \gamma \int_0^t \beta \frac{(N - I(u))I(u)}{N} e^{-\gamma(t-u)} du \\ &\quad + \beta \frac{(N - I(t))I(t)}{N} \\ &= -\gamma \left( I_0(0)e^{-\gamma t} + \int_0^t \beta \frac{(N - I(u))I(u)}{N} e^{-\gamma(t-u)} du \right) \\ &\quad + \beta \frac{(N - I(t))I(t)}{N} \\ &= \beta \frac{(N - I(t))I(t)}{N} - \gamma I(t) \end{aligned}$$

This is the classical logistic type ordinary differential equation (ODE) for  $I$  in an SIS model without vital dynamics (no birth or death)

## Case of a step function survival function

Consider case where the time spent infected has survival function

$$P(t) = \begin{cases} 1, & 0 \leq t \leq \omega, \\ 0, & t > \omega. \end{cases}$$

i.e., the sojourn time in the infective state is a constant  $\omega > 0$

In this case (8) becomes

$$I(t) = I_0(t) + \int_{t-\omega}^t \beta \frac{(N - I(u))I(u)}{N} du. \quad (10)$$

Here, it is more difficult to obtain an expression for  $I_0(t)$ . It is however assumed that  $I_0(t)$  vanishes for  $t > \omega$

When differentiated, (10) gives, for  $t \geq \omega$ ,

$$I'(t) = I'_0(t) + \beta \frac{(N - I(t))I(t)}{N} - \beta \frac{(N - I(t - \omega))I(t - \omega)}{N}.$$

Since  $I_0(t)$  vanishes for  $t > \omega$ , this gives the delay differential equation (DDE)

$$I'(t) = \beta \frac{(N - I(t))I(t)}{N} - \beta \frac{(N - I(t - \omega))I(t - \omega)}{N}.$$

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## AN EPIDEMIOLOGY MODEL THAT INCLUDES A LEAKY VACCINE WITH A GENERAL WANING FUNCTION

JULIEN ARINO

Department of Mathematics and Statistics  
University of Victoria, Victoria B.C., Canada V8W 3P4

K.L. COOKE

Department of Mathematics  
Pomona College, Claremont, CA 91711-6348 USA

P. VAN DEN DRIESSCHE

Department of Mathematics and Statistics  
University of Victoria, Victoria B.C., Canada V8W 3P4

J. VELASCO-HERNÁNDEZ

Programa de Matemáticas Aplicadas y Computación  
Instituto Mexicano del Petróleo, Eje Central Lázaro Cárdenas 152  
San Bartolo Atepehuacan, D.F. 07730, Mexico

(Communicated by Linda Allen)

## Sojourn times in compartments

### A model with vaccination

- The general model

- Case reducing to an ODE

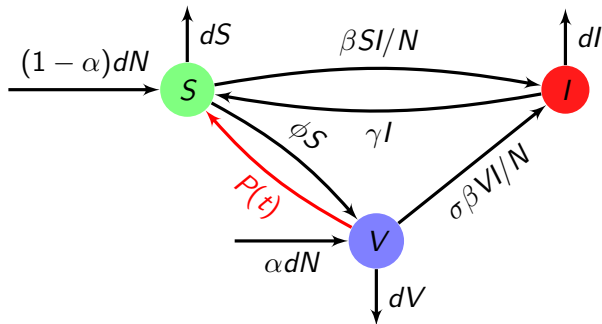
- Case reducing to a DDE

- DTMC SIS system

## A model with vaccine efficacy and waning

- ▶ Exponential distribution of recovery times (rate  $\gamma$ )
- ▶ Susceptible individuals are vaccinated (number of vaccinated at time  $t$  is denoted  $V(t)$ )
- ▶ Vaccination wanes, a fraction  $P(t)$  of the vaccinated at time  $t = 0$  remain protected by the vaccine
- ▶ Vaccination is imperfect,  $0 \leq 1 - \sigma \leq 1$  is the vaccine **efficacy**

# Model structure





# Parameters

- ▶  $d > 0$ : mortality rate
- ▶  $\gamma \geq 0$ : recovery rate
- ▶  $\beta > 0$ : infectiousness of the disease
- ▶  $\phi \geq 0$ : vaccination rate of susceptible individuals
- ▶  $\alpha \in [0, 1)$ : fraction of newborns vaccinates
- ▶  $0 \leq 1 - \sigma \leq 1$ : efficacy of the vaccine. From now on, assume  $0 \leq \sigma < 1$

- ▶ Disease transmission: standard incidence
- ▶ Vaccination of newborns
- ▶ Birth and death rate equal ( $\Rightarrow$ constant total population)

**Assumptions on  $P$ :**  $P(t)$  is a nonnegative and nonincreasing function with  $P(0^+) = 1$ , and such that  $\int_0^\infty P(u)du$  is positive and finite

Constant total population  $\Rightarrow S(t) = N - I(t) - V(t)$ ; further, we switch to **proportions**:  $S$ ,  $I$  and  $V$  represent the proportions in the population, and  $N = 1$  ( $S$  used in equations for conciseness)

## The SIS model with vaccination

$$\frac{dI(t)}{dt} = \beta(S(t) + \sigma V(t))I(t) - (d + \gamma)I(t) \quad (11a)$$

$$V(t) = V_0(t) \quad (11b)$$
$$+ \int_0^t (\phi S(u) + \alpha d)P(t-u)e^{-d(t-u)}e^{-\sigma\beta \int_u^t I(x)dx} du$$

- ▶  $\alpha d$  proportion of vaccinated newborns
- ▶  $\phi S(u)$  proportion of vaccinated susceptibles
- ▶  $P(t-u)$  fraction of the proportion vaccinated still in the  $V$  class  $t-u$  time units after going in
- ▶  $e^{-d(t-u)}$  fraction of the proportion vaccinated not dead due to natural causes
- ▶  $e^{-\sigma\beta \int_u^t I(x)dx}$  fraction of the proportion vaccinated not gone to the infective class

## Obtaining the initial condition

Let  $v(t, \tau)$  be the (density) proportion of individuals in vaccination class-age  $\tau$  still vaccinated at time  $t$ , then

$$\left( \frac{\partial}{\partial t} + \frac{\partial}{\partial \tau} \right) v(t, \tau) = -(\sigma\beta I(t) + d + \eta(\tau))v(t, \tau) \quad (12)$$

where  $V(t) = \int_0^\infty v(t, \tau) d\tau$ .  $\eta(\tau)$  is the vaccine waning rate coefficient, with proportion still in the vaccination class-age  $\tau$  being  $P(\tau) = \exp(-\int_0^\tau \eta(q) dq)$ . It is assumed that  $P$  is a survival function

Inflow in class-age zero is

$$v(t, 0) = \phi S(t) + \alpha d$$

and  $v(0, \tau) \geq 0$  is assumed

Integrating (12) along characteristics, dividing the integral for  $V(t)$  at  $t$ , substituting in the solutions, and changing integration variables, we get

$$V_0(t) = e^{-\int_0^t (\sigma\beta I(x) + d) dx} \int_0^\infty v(0, u) \frac{P(t+u)}{P(u)} du \quad (13)$$

The ratio  $P(t+u)/P(u) = \exp\left(\int_u^{t+u} \eta(q) dq\right)$  is well defined for  $t+u \geq u \geq 0$  and bounded above by 1

Since  $V(0)$  is finite, the integral in  $V_0(t)$  converges, and thus  $V_0(t)$  is nonnegative, nonincreasing and  $\lim_{t \rightarrow \infty} V_0(t) = 0$

Let

$$\mathcal{D} = \{(S, I, V); S \geq 0, I \geq 0, V \geq 0, S + I + V = 1\}$$

## Theorem 2

*The set  $\mathcal{D}$  is positively invariant under the flow of (11) with  $I(0) > 0, S(0) > 0$*

With the assumed initial conditions in  $\mathcal{D}$ , it can be shown that the system defined by (11a) and (11b) is equivalent to the system defined by (11a) and

$$\begin{aligned} \frac{d}{dt}V(t) = & \frac{d}{dt}V_0(t) + \phi S(t) + \alpha d \\ & - (d + \sigma\beta I(t))(V(t) - V_0(t)) + Q(t) \end{aligned} \quad (14)$$

where to simplify notation, we denote

$$Q(t) = \int_0^t (\phi S(u) + \alpha d) d_t(P(t-u)) e^{-d(t-u)} e^{-\sigma\beta \int_u^t I(x) dx} du$$

The system defined by (11a) and (14) is of standard form, therefore results of Hale (see Hale & Verduyn-Lunel) ensure the local existence, uniqueness and continuation of solutions of model (11)

$\mathcal{R}_0$ 

Define  $\mathcal{R}_0$  with vaccination as

$$\mathcal{R}_v = \mathcal{R}_0 \left[ \frac{1 + \sigma\phi\tilde{P} - (1 - \sigma)\alpha d\tilde{P}}{1 + \phi\tilde{P}} \right] \quad (15)$$

where  $\mathcal{R}_0 = \frac{\beta}{d+\gamma}$  is the reproduction number in the absence of vaccination and

$$\tilde{P} = \lim_{t \rightarrow \infty} \int_0^t P(v) e^{-dv} dv$$

in such a way that  $\tilde{P} < 1/d$

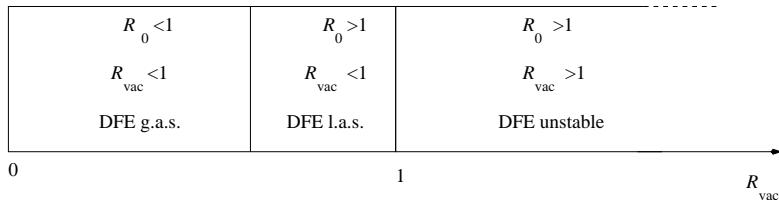
►  $\mathcal{R}_v \leq \mathcal{R}_0$  and, in absence of vaccination,  $\mathcal{R}_v = \mathcal{R}_0$



### Theorem 3

System (11) with an arbitrary loss of vaccination function  $P(t)$  always admits the disease-free equilibrium

- ▶ If  $\mathcal{R}_0 < 1$ , then the DFE is the only equilibrium of the system and the disease goes extinct
- ▶ If  $\mathcal{R}_v < 1$ , the DFE is LAS; if  $\mathcal{R}_v > 1$ , the DFE is unstable



## Sojourn times in compartments

### A model with vaccination

- The general model

- Case reducing to an ODE

- Case reducing to a DDE

- DTMC SIS system

## Reduction of the system using specific $P(t)$ functions

As before, two examples

- ▶ The distribution of waning times is exponential, which leads to an ODE system. Treated briefly here, just so as to emphasize the presence of a so-called *backward bifurcation*, a rather uncommon phenomenon in epidemiological models
  
- ▶ The waning time is a constant, which leads to a DDE model. We show that the backward bifurcation is also present

## Case reducing to an ODE system

Assume  $P(v) = e^{-\theta v}$ ,  $\theta > 0$ .  $V_0(t) = V_0(0)e^{-(d+\theta)t}e^{-\int_0^t \sigma\beta I(x)dx}$  from (13). Then (11a) and (14) give the ODE system

$$\frac{dI}{dt} = \beta(1 - I - (1 - \sigma)V)I - (d + \gamma)I \quad (16a)$$

$$\frac{dV}{dt} = \phi(1 - I - V) - \sigma\beta IV - (d + \theta)V + \alpha d \quad (16b)$$

which with no newborn vaccination ( $\alpha = 0$ ) is the model studied in Kribs-Zaletta & Velasco-Hernandez, 2000

From Theorem 3 the DFE always exists, with

$$I_{DFE} = 0, S_{DFE} = \frac{\theta + d(1 - \alpha)}{d + \theta + \phi}, V_{DFE} = \frac{\phi + \alpha d}{d + \theta + \phi}$$

## Backward bifurcation

Assume that  $\mathcal{R}_0 > 1$ , then endemic equilibria (positive  $I$  equilibria, denoted by  $I^*$ ) can be obtained analytically from the quadratic equation

$$\mathcal{P}(I) = AI^2 + BI + C = 0$$

where

$$A = -\sigma\beta$$

$$B = \sigma(\beta - (d + \gamma)) - (d + \theta + \sigma\phi)$$

$$C = (d + \gamma)(d + \theta + \phi)(\mathcal{R}_v - 1)/\beta$$

with

$$\mathcal{R}_v = \mathcal{R}_0 \frac{d + \theta + \sigma\phi - \alpha(1 - \sigma)d}{d + \theta + \phi}$$

from (15).

Backward bifurcation leading to two endemic equilibria occurs for  $\sigma > 0$  if  $\mathcal{P}'(0) = B > 0$ ,  $\mathcal{P}(0) = C < 0$  and  $B^2 > 4AC$  (we always have  $\mathcal{P}(1) < 0$ )

► On an  $(\mathcal{R}_v, I)$  bifurcation diagram, this occurs for  $\mathcal{R}_c < \mathcal{R}_v < 1$ , where  $\mathcal{R}_c$  is the value of  $\mathcal{R}_v$  at the saddle node bifurcation point where the two values of  $I$  coincide, *i.e.*,  $I = I_c = B/(-2A)$

► For  $\mathcal{R}_v < \mathcal{R}_c$ , there is no endemic equilibrium (EEP). For  $\mathcal{R}_v > 1$ , the constant term  $C > 0$ , and there is a unique EEP

► In the case of forward bifurcation,  $\mathcal{R}_c = 1$ ; this is the case in particular if the vaccine is totally effective ( $\sigma = 0$ )

By standard planar ODE arguments the following can be shown

#### Theorem 4

*For the ODE system (16) with  $V(0) \geq 0$ ,  $I(0) > 0$ , and  $\mathcal{R}_0 > 1$*

*(i) if  $\mathcal{R}_v < \mathcal{R}_c$ , then the disease dies out*

*(ii) if  $\mathcal{R}_c < \mathcal{R}_v < 1$ , then the EEP with larger  $I$  is l.a.s., and the EEP with smaller  $I$  is unstable*

*(iii) if  $\mathcal{R}_v > 1$ , then the unique EEP is globally asymptotically stable in  $\mathcal{D} \setminus \{I = 0\}$*

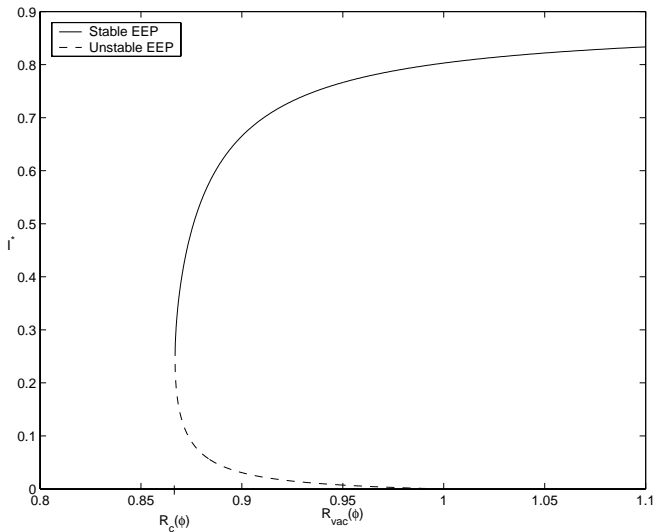
## Pertussis:

- ▶ 3 week average disease duration ( $\gamma = 0.04762$ )
- ▶ Average lifetime 75 years ( $d = 3.6530E - 05$ )
- ▶ Average number of adequate contacts per infective per day is estimated at 0.4 ( $\beta = 0.4$ )
- ▶ Most newborns are vaccinated in the first few months of life ( $\alpha = 0.9$ )
- ▶ Vaccine is effective,  $\sigma = 0.1$  (90% effective vaccine).
- ▶ Pertussis vaccine begins to wane after about 3 years and the average waning time of the vaccine  $1/\theta$  is assumed to be 5 years, giving  $\theta = 5.4794E - 04$

With these parameter values, there is backward bifurcation for a range of  $\phi$  values given by  $0.0254 \leq \phi \leq 0.1506$



With the above parameter values,  $\mathcal{R}_0 = 8.3936$  and  $\mathcal{R}_v(\phi) = 0.8807$  for  $\phi = 0.1$ , which is in the range of backward bifurcation since the critical value  $\mathcal{R}_c(\phi) = 0.8669 < \mathcal{R}_v(\phi) < 1$



## Sojourn times in compartments

### A model with vaccination

- The general model

- Case reducing to an ODE

- Case reducing to a DDE

- DTMC SIS system

## Step function case: a delay integral model

Suppose that

$$P(v) = \begin{cases} 1 & \text{if } v \in [0, \omega] \\ 0 & \text{otherwise} \end{cases}$$

Since  $V_0(t) = 0$  for  $t > \omega$ , with  $S = 1 - I - V$  the integral equation (11b) becomes, for  $t > \omega$

$$V(t) = \int_{t-\omega}^t (\phi(1 - I(u) - V(u)) + \alpha d) e^{-d(t-u)} e^{-\sigma\beta \int_u^t I(x) dx} du \quad (17)$$

Differentiating (17) (see equation (14)) gives the model as the two dimensional system, for  $t > \omega$

$$\frac{d}{dt}I(t) = \beta(1 - I(t) - (1 - \sigma)V(t))I(t) - (d + \gamma)I(t) \quad (18a)$$

$$\frac{d}{dt}V(t) = \phi(1 - I(t) - V(t)) \quad (18b)$$

$$\begin{aligned} & - \phi(1 - I(t - \omega) - V(t - \omega))e^{-d\omega} e^{-\sigma\beta \int_{t-\omega}^t I(x)dx} \\ & - \sigma\beta IV - dV + \alpha d \left( 1 - e^{-d\omega} e^{-\sigma\beta \int_{t-\omega}^t I(x)dx} \right) \end{aligned}$$

Hereafter, shift time by  $\omega$  so that these equations hold for  $t > 0$

The well posedness of the problem follows from Theorem 2 and from the fact that solutions of (11) exist and are unique. For a constant waning period, the basic reproduction number from (15) is

$$\mathcal{R}_v = \mathcal{R}_0 \frac{d + (\sigma\phi - \alpha(1 - \sigma)d)(1 - e^{-d\omega})}{d + \phi(1 - e^{-d\omega})} \quad (19)$$

With  $I_{DF} = 0$ , from Theorem 3

$$V_{DF} = \frac{(\phi + \alpha d)(1 - e^{-d\omega})}{d + \phi(1 - e^{-d\omega})}, \quad S_{DF} = \frac{d - \alpha d(1 - e^{-d\omega})}{d + \phi(1 - e^{-d\omega})} \quad (20)$$

## Finding the EEP's

From nullclines, there exists one (or more) endemic equilibria (EEP) iff there exists  $0 < I^* \leq 1$  such that

$$V^* = f(I^*) = g(I^*) \quad (21)$$

where

$$f(I) = \frac{1 - 1/\mathcal{R}_0 - I}{1 - \sigma} \quad (22)$$

for  $\sigma < 1$ , and

$$g(I) = \frac{(\phi(1 - I) + \alpha d)(1 - e^{-d\omega - \sigma\beta\omega I})}{\phi(1 - e^{-d\omega - \sigma\beta\omega I}) + d + \sigma\beta I} \quad (23)$$

## Visualising and locating the bifurcation

From the nullcline equations, an EEP exists iff there exists an  $I^* \in (0, 1]$  such that equations (21)-(23) hold. So we study the zeros of

$$H(I) = \frac{1 - 1/\mathcal{R}_0 - I}{1 - \sigma} - \frac{(\phi(1 - I) + \alpha d)(1 - e^{-d\omega - \sigma\beta\omega I})}{\phi(1 - e^{-d\omega - \sigma\beta\omega I}) + d + \sigma\beta I}$$

To state the problem in a formal way, let  $\mathcal{A} = \{\alpha, \beta, \gamma, \omega, \phi, \sigma\}$  be the set of parameters of interest, and denote

$$H(I, \mathcal{A}) = f(I) - g(I) \tag{24}$$

to show the dependence on these parameters.

Proceed as follows

1. Choose a parameter  $a_i \in \mathcal{A}$
2. Fix all other  $a_j$ 's ( $j \neq i$ )
3. Choose  $a_{i,min}$ ,  $a_{i,max}$  and  $\Delta a_i$  for  $a_i$
4. For all  $a_{i,k} = a_{i,min} + k\Delta a_i$  ( $k$  such that  $a_{i,k} \leq a_{i,max}$ ), compute  $I^*$  such that  $H(I^*, a_{i,k}) = 0$

Step 4 is carried out using the MATLAB `fzero` function

Further precision can be gained by showing that

$$H(0) = \frac{\mathcal{R}_v - 1}{(1 - \sigma)\mathcal{R}_0}$$

and that, for  $\sigma < 1$

$$H(1) = -\frac{1}{(1 - \sigma)\mathcal{R}_0} - \frac{\alpha d(1 - e^{-d\omega - \sigma\beta\omega})}{\phi(1 - e^{-d\omega - \sigma\beta\omega}) + d + \sigma\beta} < 0$$



Define  $\mathcal{R}_c$  as previously. For  $\mathcal{R}_0 > 1$  and  $\mathcal{R}_v < 1$ , there are several possibilities

- ▶ If  $\mathcal{R}_v < \mathcal{R}_c$ , then there is no EEP.  $H(0)$  and  $H(1)$  are strictly negative, and numerical simulations seem to indicate that  $H$  has no roots in  $(0, 1]$  (i.e., that  $H < 0$  on this interval)
- ▶ If  $\mathcal{R}_c < \mathcal{R}_v < 1$ , then there are endemic equilibria. Here, since  $H(0)$  and  $H(1)$  are strictly negative, the only possibility is thus to have an even number of zeros of  $H$ . Numerical simulations appear to indicate that the number of endemic equilibria is 2

In between these two situations  $\mathcal{R}_v = \mathcal{R}_c$  and there is one endemic equilibrium  $I^*$ . Using the same procedure as for the visualisation of the bifurcation, it is possible to compute  $\mathcal{R}_c$  by finding the value  $I^*$  such that  $H(I^*, \mathcal{A}) = 0$  and  $H'(I^*, \mathcal{A}) = 0$ , for a given parameter  $a_i \in \mathcal{A}$

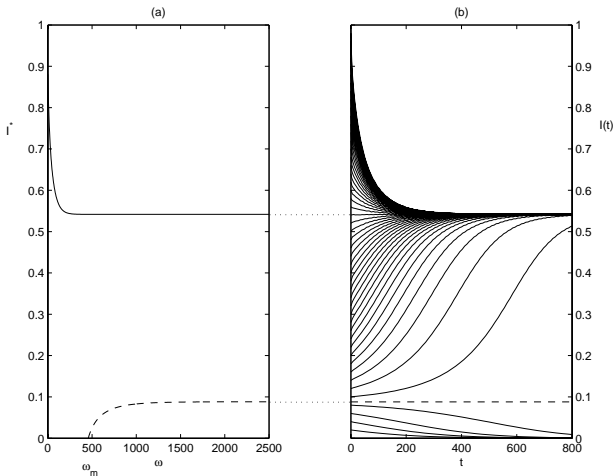
If  $\mathcal{R}_v > 1$  then  $H(0) > 0$  and so there is an odd number of endemic equilibria. Numerical simulations indicate that there is a unique EEP

## Numerical bifurcation analysis

Same parameter values as in ODE case, except that the constant waning time (the delay)  $\omega$  has to be substituted for  $\theta$ . We take  $\omega = 1825$ , *i.e.*, corresponding to a 5 years waning time

These parameters give  $\mathcal{R}_0 = 8.3936$  and  $\mathcal{R}_v(\phi) = 0.8819$ , which is in the range of the backward bifurcation since (using the above method)  $\mathcal{R}_c(\phi) = 0.8675$

The bifurcation diagram is very like that depicted in earlier for the ODE. Numerical simulations of the DDE model (using `dde23`) indicate that there are no additional bifurcations; solutions either go to the DFE or to the (larger) EEP



(a) Values of  $I^*$  as a function of  $\omega$  by solving  $H(I, \mathcal{A}) = 0$  with  $a_i = \omega$ . (b) Value of  $I(t)$  versus time, obtained by numerical integration of system (18) with initial data  $I(t) = c$ , for  $t \in [-\omega, 0]$ ,  $\omega = 1825$ ,  $c$  varying from 0 to 1 by steps of 0.02

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## Age of vaccination

We have seen that infinite dimensionality can result from a detailed description (or an unspecified one) of the sojourn time in compartments

We used age of vaccination to find the initial condition of (11)

Here we take a closer look at this type of model

Originally, age of infection was introduced to account for differences in infectivity depending on the time since an individual became infected

For instance, it is known that infectiousness of HIV positive patients vary as a function of time

## EVALUATION OF VACCINATION STRATEGIES DURING PANDEMIC OUTBREAKS

CHRISTOPHER S. BOWMAN

Institute for Biodiagnostics  
National Research Council Canada  
435 Ellice Avenue, Winnipeg, MB, R3B 1Y6, Canada

JULIEN ARINO

Department of Mathematics  
University of Manitoba  
Winnipeg, Manitoba, R3T 2N2, Canada

SEYED M. MOGHADAS

Institute for Biodiagnostics  
National Research Council Canada  
435 Ellice Avenue, Winnipeg, MB, R3B 1Y6, Canada  
and

Department of Mathematics and Statistics  
University of Winnipeg  
Winnipeg, MB, R3B 2E9, Canada

## How to model time between vaccine doses

$$S' = -fS - V_1(t, 0) \quad (25a)$$

$$A' = \left( (1-p)S + (1-p_1)\delta_1 \tilde{V}_1 + (1-p_2)\delta_2 V_2 \right) f - \mu_A A \quad (25b)$$

$$I' = (pS + p_1\delta_1 \tilde{V}_1 + p_2\delta_2 V_2) f - \mu I \quad (25c)$$

$$V_2' = V_1(t, a^*) - \delta_2 f V_2(t) \quad (25d)$$

$$\left( \frac{\partial}{\partial t} + \frac{\partial}{\partial a} \right) V_1(t, a) = -\delta_1 f V_1(t, a), \quad 0 \leq a \leq a^* \quad (25e)$$

and boundary condition

$$V_1(t, 0) = \begin{cases} \gamma S_0 \left( \frac{S(t)}{S(t)+A(t)} \right) & \text{if } T \leq t \leq T_e \text{ and } S > 0 \\ 0 & \text{otherwise} \end{cases} \quad (25f)$$

where  $f = \beta(\delta_A A + I)$  and  $\tilde{V}_1(t) = \int_0^{a^*} V_1(t, a) da$

## Simplifying a bit

Integrate (25e) using characteristics along lines  $a = s$  and  $t = T + s$ , with  $s$  as a new variable

$$V_1(t, a) = V_1(t - a, 0) \exp \left( \int_{t-a}^t -\delta_1 f(\xi) d\xi \right) \quad (26)$$

Define

$$\zeta(t) = \int_0^t \delta_1 f(\xi) d\xi$$

and substitute into (26), giving

$$V_1(t, a) = V_1(t - a, 0) \exp (\zeta(t - a)\zeta(t))$$

So the distributed delay is now discrete



## Simplifying a bit more

Let

$$\nu(t) = \int_0^t V_1(s, 0) e^{\zeta(s)} ds$$

Then the total number of individuals having been vaccinated with a single dose is

$$\tilde{V}_1(t) = e^{-\zeta(t)} (\nu(t) - \nu(t - a^*))$$

$$S' = -fS - V_1(t, 0) \quad (27a)$$

$$A' = \left( (1 - p)S + (1 - p_1)\delta_1 \tilde{V}_1 + (1 - p_2)\delta_2 V_2 \right) f - \mu_A A \quad (27b)$$

$$I' = (pS + p_1\delta_1 \tilde{V}_1 + p_2\delta_2 V_2) f - \mu I \quad (27c)$$

$$V_2' = V_1(t - a^*, 0) e^{\zeta(t - a^*)} - \delta_2 f V_2(t) \quad (27d)$$

$$\zeta' = \delta_1 f \quad (27e)$$

$$\nu' = V_1(t, 0) e^{\zeta(t)} \quad (27f)$$

## Sojourn times in compartments

Distributions of times to events

Two “extreme” distributions and a nicer one

A simple cohort model with death

A possible fix to the exponential distribution issue

Sojourn times in an SIS disease transmission model

A model with vaccination

Age of vaccination

Conclusions on sojourn times

## Conclusions on sojourn times

- ▶ The time of sojourn in compartments plays an important role in determining the type of model that we deal with
- ▶ All ODE models, when they use terms of the form  $\kappa X$ , make the assumption that the time of sojourn in compartments is exponentially distributed with parameter  $\kappa$
- ▶ At the other end of the spectrum, delay differential with discrete delay make the assumption of a constant sojourn time, equal for all individuals
- ▶ Both can be true sometimes... but reality is more likely somewhere in between

Sojourn times in compartments

Discrete-time Markov chains

Continuous time Markov chains

Structuration in age

A discrete-time Markov chain takes the form

$$p(n+1) = p(n)P, \quad n = 1, 2, 3, \dots$$

where  $p(n) = (p_1(n), p_2(n), \dots, p_r(n))$  is a (row) probability vector and  $P = (p_{ij})$  is a  $r \times r$  **transition matrix**

$$P = \begin{pmatrix} p_{11} & p_{12} & \cdots & p_{1r} \\ p_{21} & p_{22} & \cdots & p_{2r} \\ p_{r1} & p_{r2} & \cdots & p_{rr} \end{pmatrix}$$

# Stochastic matrices

## Definition 5

The nonnegative  $r \times r$  matrix  $M$  is (**row**) **stochastic** if  $\sum_{j=1}^r a_{ij} = 1$  for  $i = 1, 2, \dots, r$

## Definition 6

Let  $M$  be a stochastic matrix  $M$ . Then all eigenvalues  $\lambda$  of  $M$  are such that  $|\lambda| \leq 1$ . Furthermore,  $\lambda = 1$  is an eigenvalue of  $M$

## Theorem 7

*If  $M, N$  are stochastic matrices, then  $MN$  is a stochastic matrix*

## Theorem 8

*If  $M$  is a stochastic matrix, then for any  $k \in \mathbb{N}$ ,  $M^k$  is a stochastic matrix*

## Asymptotic behavior

Let  $p(0)$  be the initial distribution (row) vector. Then

$$\begin{aligned}p(1) &= p(0)P \\p(2) &= p(1)P \\&= (p(0)P)P \\&= p(0)P^2\end{aligned}$$

Iterating, we get that for any  $n$ ,

$$p(n) = p(0)P^n$$

Therefore,

$$\lim_{n \rightarrow +\infty} p(n) = \lim_{n \rightarrow +\infty} p(0)P^n = p(0) \lim_{n \rightarrow +\infty} P^n$$

## Discrete-time Markov chains

Regular DTMC

Random walk v1.0 (regular case)

Absorbing DTMC



# Regular Markov chain

## Definition 9

A **regular Markov chain** is one in which  $P^k$  is positive for some integer  $k > 0$ , i.e.,  $P^k$  has only positive entries, no zero entries

## Definition 10

A nonnegative matrix  $M$  is **primitive** if, and only if, there is an integer  $k > 0$  such that  $M^k$  is positive

## Theorem 11

*A Markov chain is regular if, and only if, the transition matrix  $P$  is primitive*

## Important result for regular Markov chains

### Theorem 12

If  $P$  is the transition matrix of a regular Markov chain, then

1. the powers  $P^n$  approach a stochastic matrix  $W$
2. each row of  $W$  is the same (row) vector  $w = (w_1, \dots, w_r)$
3. the components of  $w$  are positive

So if the Markov chain is regular

$$\lim_{n \rightarrow +\infty} p(n) = p(0) \lim_{n \rightarrow +\infty} P^n = p(0)W$$

The vector  $w$  is the left eigenvector corresponding to the eigenvalue 1 of  $P$ . (We already know that the (right) eigenvector corresponding to 1 is  $\mathbb{1}$ .)

Indeed, if  $p(n)$  converges, then  $p(n+1) = p(n)P$ , so  $w$  is a fixed point of the system. We thus write

$$wP = w$$

and solve for  $w$ , which amounts to finding  $w$  as the left eigenvector corresponding to the eigenvalue 1

Alternatively, we can find  $w$  as the (right) eigenvector associated to the eigenvalue 1 for the transpose of  $P$

$$P^T w^T = w^T$$

(normalise if need be)

# Linking matrix and graph theory

## Definition 13

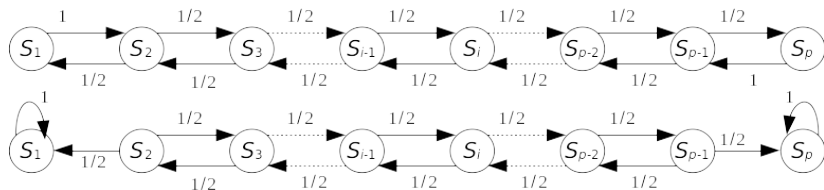
A digraph  $\mathcal{G}$  is **strongly connected** if there is a path between all pairs of vertices

## Definition 14

A matrix  $M \in \mathcal{M}_n$  is **irreducible** if there does not exist a matrix  $P \in \mathcal{M}_n$  s.t.  $P^{-1}AP$  block triangular

## Theorem 15

$A \in \mathcal{M}_n$  irreducible  $\iff \mathcal{G}(A)$  strongly connected



## Discrete-time Markov chains

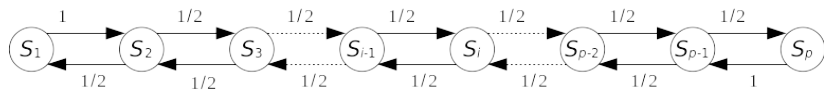
Regular DTMC

Random walk v1.0 (regular case)

Absorbing DTMC

## Random walk 1.0 (regular case)

- ▶ chain of states  $S_1, \dots, S_p$
- ▶ if in state  $S_i$ ,  $i = 2, \dots, p - 1$ , probability  $1/2$  of going left (to  $S_{i-1}$ ) and  $1/2$  of going right (to  $S_{i+1}$ )
- ▶ if in state  $S_1$ , probability 1 of going to  $S_2$
- ▶ if in state  $S_p$ , probability 1 of going to  $S_{p-1}$



## Transition matrix for RW 1.0

$$P = \begin{pmatrix} 0 & 1 & 0 & 0 & 0 & \cdots & 0 \\ 1/2 & 0 & 1/2 & 0 & & & \\ 0 & 1/2 & 0 & 1/2 & & & \\ \vdots & & \ddots & \ddots & \ddots & & \vdots \\ & & & & & 1/2 & 0 & 1/2 \\ & & & & & 0 & 1 & 0 \end{pmatrix}$$

Clearly a primitive matrix, so a regular Markov chain. We find (easy to do by hand)

$$w^T = \left( \frac{1}{2(p-1)}, \frac{1}{p-1}, \dots, \frac{1}{p-1}, \frac{1}{2(p-1)} \right)$$

## Setting up the transition matrix

```
# Total population
nb_states = 10 # Small so we can see output
# Parameters
proba_left = 0.5
proba_right = 0.5
proba_stay = 1-(proba_left+proba_right)
# Make the transition matrix
T = mat.or.vec(nr = nb_states, nc = nb_states)
for (row in 2:(nb_states-1)) {
  T[row,(row-1)] = proba_left
  T[row,(row+1)] = proba_right
  T[row, row] = proba_stay
}
# First row only has move right
T[1,2] = 1
# Last row only has move left
T[nb_states, (nb_states-1)] = 1
```



## Analysis using markovchain library

```
library(markovchain)
mcRW <- new("markovchain",
           states = sprintf("S_%d", 1:nb_states),
           transitionMatrix = T,
           name = "RW_reg")
```

```
> summary(mcRW)
RW_reg Markov chain that is composed by:
Closed classes:
S_1 S_2 S_3 S_4 S_5 S_6 S_7 S_8 S_9 S_10
Recurrent classes:
{S_1,S_2,S_3,S_4,S_5,S_6,S_7,S_8,S_9,S_10}
Transient classes:
NONE
The Markov chain is irreducible
The absorbing states are: NONE
```

```

> steadyStates(mcrW)
      S_1      S_2      S_3      S_4      S_5      S_6
      S_7      S_8      S_9
[1,] 0.05555556 0.11111111 0.11111111 0.11111111 0.11111111 0.11111111
      0.11111111 0.11111111 0.11111111
      S_10
[1,] 0.05555556

```

Jives with

$$w^T = \left( \frac{1}{2(p-1)}, \frac{1}{p-1}, \dots, \frac{1}{p-1}, \frac{1}{2(p-1)} \right)$$

we had computed

`meanRecurrenceTime`: outputs a named vector with the expected time to first return to a state when the chain starts there. States present in the vector are only the recurrent ones. If the matrix is ergodic (i.e. irreducible), then all states are present in the output and order is the same as states order for the Markov chain

```
> meanRecurrenceTime(mcRW)
 S_1 S_2 S_3 S_4 S_5 S_6 S_7 S_8 S_9 S_10
 18  9  9  9  9  9  9  9  9  18
```

`period`: returns a integer number corresponding to the periodicity of the Markov chain (if it is irreducible)

```
> period(mcRW)
[1] 2
```

(period of state  $x \in \mathcal{S}$  is  $\gcd\{n \in \mathbb{N}_+ : T^n(x, x) > 0\}$ )

meanFirstPassageTime: Given an irreducible (ergodic) markovchain object, this function calculates the expected number of steps to reach other states

```
> meanFirstPassageTime(mcRW)
```

	S_1	S_2	S_3	S_4	S_5	S_6	S_7	S_8	S_9	S_10
S_1	0	1	4	9	16	25	36	49	64	81
S_2	17	0	3	8	15	24	35	48	63	80
S_3	32	15	0	5	12	21	32	45	60	77
S_4	45	28	13	0	7	16	27	40	55	72
S_5	56	39	24	11	0	9	20	33	48	65
S_6	65	48	33	20	9	0	11	24	39	56
S_7	72	55	40	27	16	7	0	13	28	45
S_8	77	60	45	32	21	12	5	0	15	32
S_9	80	63	48	35	24	15	8	3	0	17
S_10	81	64	49	36	25	16	9	4	1	0

## Discrete-time Markov chains

Regular DTMC

Random walk v1.0 (regular case)

Absorbing DTMC

## Absorbing states, absorbing chains

### Definition 16

A state  $S_i$  in a Markov chain is **absorbing** if whenever it occurs on the  $n^{\text{th}}$  generation of the experiment, it then occurs on every subsequent step. In other words,  $S_i$  is absorbing if  $p_{ii} = 1$  and  $p_{ij} = 0$  for  $i \neq j$

### Definition 17

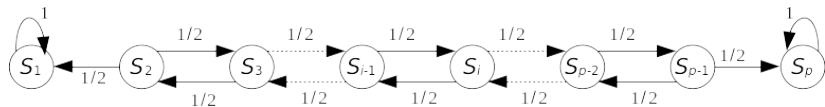
A **Markov chain is absorbing** if it has at least one absorbing state, and if from every state it is possible to go to an absorbing state

### Definition 18

In an absorbing Markov chain, a state that is not absorbing is called **transient**

## Some questions on absorbing chains

Suppose we have a chain like the following



1. Does the process eventually reach an absorbing state?
2. Average number of times spent in a transient state, if starting in a transient state?
3. Average number of steps before entering an absorbing state?
4. Probability of being absorbed by a given absorbing state, when there are more than one, when starting in a given transient state?

# Reaching an absorbing state

Answer to question 1:

## Theorem 19

*In an absorbing Markov chain, the probability of reaching an absorbing state is 1*



## Standard form of the transition matrix

For an absorbing chain with  $k$  absorbing states and  $r - k$  transient states, the transition matrix can be written as

$$P = \begin{pmatrix} \mathbb{I}_k & \mathbf{0} \\ R & Q \end{pmatrix}$$

	Absorbing states	Transient states
Absorbing states	$\mathbb{I}_k$	$\mathbf{0}$
Transient states	$R$	$Q$

$\mathbb{I}_k$  the  $k \times k$  identity,  $\mathbf{0} \in \mathbb{R}^{k \times (r-k)}$ ,  $R \in \mathbb{R}^{(r-k) \times k}$ ,  $Q \in \mathbb{R}^{(r-k) \times (r-k)}$

The matrix  $\mathbb{I}_{r-k} - Q$  is invertible. Let

- ▶  $N = (\mathbb{I}_{r-k} - Q)^{-1}$  be the **fundamental matrix** of the Markov chain
- ▶  $T_i$  be the sum of the entries on row  $i$  of  $N$
- ▶  $B = NR$

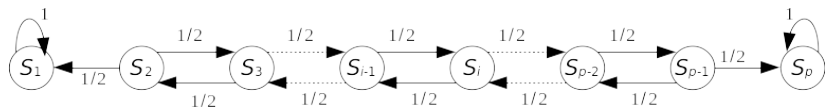
Answers to our remaining questions:

2.  $N_{ij}$  is the average number of times the process is in the  $j$ th transient state if it starts in the  $i$ th transient state
3.  $T_i$  is the average number of steps before the process enters an absorbing state if it starts in the  $i$ th transient state
4.  $B_{ij}$  is the probability of eventually entering the  $j$ th absorbing state if the process starts in the  $i$ th transient state

See for instance book of Kemeny and Snell

## Random walk 2.0 (absorbing case)

- ▶ chain of states  $S_1, \dots, S_p$
- ▶ if in state  $S_i$ ,  $i = 2, \dots, p - 1$ , probability  $1/2$  of going left (to  $S_{i-1}$ ) and  $1/2$  of going right (to  $S_{i+1}$ )
- ▶ if in state  $S_1$ , probability 1 of going to  $S_1$
- ▶ if in state  $S_p$ , probability 1 of going to  $S_p$



## Transition matrix for DMW 2.0

$$P = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 & \dots & 0 \\ 1/2 & 0 & 1/2 & 0 & & & \\ 0 & 1/2 & 0 & 1/2 & & & \\ \vdots & & \ddots & \ddots & \ddots & & \vdots \\ & & & & & 1/2 & 0 & 1/2 \\ & & & & & 0 & 0 & 1 \end{pmatrix}$$

## Put $P$ in standard form

Absorbing states are  $S_1$  and  $S_p$ , write them first, then write other states

	$S_1$	$S_p$	$S_2$	$S_3$	$S_4$	$\dots$	$S_{p-2}$	$S_{p-1}$
$S_1$	1	0	0	0	0	$\dots$	0	0
$S_p$	0	1	0	0	0	$\dots$	0	0
$S_2$	1/2	0	0	1/2	0	$\dots$	0	0
$S_3$	0	0	1/2	0	1/2	$\dots$	0	0
$\vdots$								
$S_{p-2}$	0	0	0	0	0	$\dots$	0	1/2
$S_{p-1}$	0	1/2	0	0	0	$\dots$	1/2	0

So we find

$$P = \begin{pmatrix} \mathbb{I}_2 & \mathbf{0} \\ R & Q \end{pmatrix}$$

where  $\mathbf{0}$  a  $2 \times (p - 2)$ -matrix,  $R$  a  $(p - 2) \times 2$  matrix and  $Q$  a  $(p - 2) \times (p - 2)$  matrix

$$R = \begin{pmatrix} 1/2 & 0 \\ 0 & 0 \\ \vdots & \vdots \\ 0 & 0 \\ 0 & 1/2 \end{pmatrix}$$

and

$$Q = \begin{pmatrix} 0 & 1/2 & 0 & & & \\ 1/2 & 0 & 1/2 & & & \\ 0 & 1/2 & 0 & & & \\ & & \ddots & \ddots & \ddots & \\ 0 & & & 1/2 & 0 & 1/2 \\ 0 & & & & 1/2 & 0 \end{pmatrix}$$

$$\mathbb{I}_{p-2} - Q = \begin{pmatrix} 1 & -1/2 & 0 & & & \\ -1/2 & 1 & -1/2 & & & \\ 0 & -1/2 & 1 & & & \\ & & & \ddots & \ddots & \ddots \\ 0 & & & & -1/2 & 1 & -1/2 \\ 0 & & & & & -1/2 & 1 \end{pmatrix}$$

This is a **symmetric tridiagonal Toeplitz** matrix

(symmetric: obvious; tridiagonal: there are three diagonal bands;  
Toeplitz: each diagonal band is constant)

Could invert it explicitly, let us not bother



## Setting up the transition matrix

```
# Total population
nb_states = 10 # Small so we see output
# Parameters
proba_left = 0.5
proba_right = 0.5
proba_stay = 1-(proba_left+proba_right)
# Make the transition matrix
T = mat.or.vec(nr = nb_states, nc = nb_states)
for (row in 2:(nb_states-1)) {
  T[row,(row-1)] = proba_left
  T[row,(row+1)] = proba_right
  T[row, row] = proba_stay
}
# First and last rows only have stay
T[1,1] = 1
T[nb_states, nb_states] = 1
```

## Analysis using markovchain library

```
library(markovchain)
mcRW <- new("markovchain",
           states = sprintf("S_%d", 1:nb_states),
           transitionMatrix = T,
           name = "RW_abs")
```

```
> summary(mcRW)
RW_abs Markov chain that is composed by:
Closed classes:
S_1
S_10
Recurrent classes:
{S_1},{S_10}
Transient classes:
{S_2,S_3,S_4,S_5,S_6,S_7,S_8,S_9}
The Markov chain is not irreducible
The absorbing states are: S_1 S_10
```

```
> canonicForm(mcRW)
```

```
RW_abs
```

A 10 - dimensional discrete Markov Chain defined by the following states:

S\_1, S\_10, S\_2, S\_3, S\_4, S\_5, S\_6, S\_7, S\_8, S\_9

The transition matrix (by rows) is defined as follows:

	S_1	S_10	S_2	S_3	S_4	S_5	S_6	S_7	S_8	S_9
S_1	1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
S_10	0.0	1.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
S_2	0.5	0.0	0.0	0.5	0.0	0.0	0.0	0.0	0.0	0.0
S_3	0.0	0.0	0.5	0.0	0.5	0.0	0.0	0.0	0.0	0.0
S_4	0.0	0.0	0.0	0.5	0.0	0.5	0.0	0.0	0.0	0.0
S_5	0.0	0.0	0.0	0.0	0.5	0.0	0.5	0.0	0.0	0.0
S_6	0.0	0.0	0.0	0.0	0.0	0.5	0.0	0.5	0.0	0.0
S_7	0.0	0.0	0.0	0.0	0.0	0.0	0.5	0.0	0.5	0.0
S_8	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.5	0.0	0.5
S_9	0.0	0.5	0.0	0.0	0.0	0.0	0.0	0.0	0.5	0.0

```

> meanAbsorptionTime(mcRW)
S_2 S_3 S_4 S_5 S_6 S_7 S_8 S_9
   8  14  18  20  20  18  14   8
> absorptionProbabilities(mcRW)
      S_1      S_10
S_2 0.8888889 0.1111111
S_3 0.7777778 0.2222222
S_4 0.6666667 0.3333333
S_5 0.5555556 0.4444444
S_6 0.4444444 0.5555556
S_7 0.3333333 0.6666667
S_8 0.2222222 0.7777778
S_9 0.1111111 0.8888889

```

hittingProbabilities: given a markovchain object, this function calculates the probability of ever arriving from state i to j

```
> hittingProbabilities(mcRW)
```

	S_1	S_2	S_3	S_4	S_5	S_6	S_7
S_1	1.0000000	0.0000	0.0000000	0.0000000	0.000	0.000	0.0000000
S_2	0.8888889	0.4375	0.5000000	0.3333333	0.250	0.200	0.1666667
S_3	0.7777778	0.8750	0.6785714	0.6666667	0.500	0.400	0.3333333
S_4	0.6666667	0.7500	0.8571429	0.7500000	0.750	0.600	0.5000000
S_5	0.5555556	0.6250	0.7142857	0.8333333	0.775	0.800	0.6666667
S_6	0.4444444	0.5000	0.5714286	0.6666667	0.800	0.775	0.8333333
S_7	0.3333333	0.3750	0.4285714	0.5000000	0.600	0.750	0.7500000
S_8	0.2222222	0.2500	0.2857143	0.3333333	0.400	0.500	0.6666667
S_9	0.1111111	0.1250	0.1428571	0.1666667	0.200	0.250	0.3333333
S_10	0.0000000	0.0000	0.0000000	0.0000000	0.000	0.000	0.0000000

The general model  
Case reducing to an ODE  
Case reducing to a DDE

## Discrete-time Markov chains

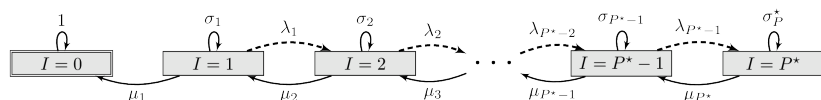
Absorbing DTMC

DTMC SIS system

# DTMC SIS system

Since  $S = P^* - I$ , consider only the infected. To simulate as DTMC, consider a random walk on  $I$  ( $\simeq$  Gambler's ruin problem)

Denote  $\lambda_I = \beta(P^* - I)\Delta t$ ,  $\mu_I = \gamma I\Delta t$  and  $\sigma_I = 1 - (\lambda_I + \mu_I)\Delta t$



To make things easy to see: Pop=5

```
# Make the transition matrix
T = mat.or.vec(nr = (Pop+1), nc = (Pop+1))
for (row in 2:Pop) {
  I = row-1
  mv_right = gamma*I*Delta_t # Recoveries
  mv_left = beta*I*(Pop-I)*Delta_t # Infections
  T[row,(row-1)] = mv_right
  T[row,(row+1)] = mv_left
}
# Last row only has move left
T[(Pop+1),Pop] = gamma*(Pop)*Delta_t
# Check that we don't have too large values
if (max(rowSums(T))>1) {
  T = T/max(rowSums(T))
}
diag(T) = 1-rowSums(T)
```



## Analysis using markovchain library

```
library(markovchain)
mcSIS <- new("markovchain",
            states = sprintf("I_%d", 0:Pop),
            transitionMatrix = T,
            name = "SIS")
```

```
> summary(mcSIS)
SIS Markov chain that is composed by:
Closed classes:
I_0
Recurrent classes:
{I_0}
Transient classes:
{I_1,I_2,I_3,I_4,I_5}
The Markov chain is not irreducible
The absorbing states are: I_0
```

```

> canonicForm(mcSIS)
SIS
  A 6 - dimensional discrete Markov Chain defined by the
  following states:
  I_0, I_1, I_2, I_3, I_4, I_5
  The transition matrix (by rows) is defined as follows:
           I_0      I_1      I_2      I_3      I_4      I_5
I_0 1.000000 0.000000 0.000000 0.000000 0.000000 0.000000
I_1 0.166667 0.500000 0.333333 0.000000 0.000000 0.000000
I_2 0.000000 0.333333 0.166667 0.500000 0.000000 0.000000
I_3 0.000000 0.000000 0.500000 0.000000 0.500000 0.000000
I_4 0.000000 0.000000 0.000000 0.666667 0.000000 0.333333
I_5 0.000000 0.000000 0.000000 0.000000 0.833333 0.166667

```

```
# The vector of steady states. Here, all mass should be in I_0
> steadyStates(mcSIS)
      I_0 I_1 I_2 I_3 I_4 I_5
[1,]  1  0  0  0  0  0
```

```
> hittingProbabilities(mcSIS)
      I_0      I_1      I_2      I_3      I_4      I_5
I_0  1 0.000000 0.000000 0.000000 0.000000 0.000000
I_1  1 0.833333 0.666667 0.545454 0.461538 0.352941
I_2  1 1.000000 0.888889 0.818181 0.692307 0.529411
I_3  1 1.000000 1.000000 0.909090 0.846153 0.647058
I_4  1 1.000000 1.000000 1.000000 0.897435 0.764705
I_5  1 1.000000 1.000000 1.000000 1.000000 0.803921
```

Read by row: if the process starts in  $I_i$  (row  $i - 1$ ), probability that state  $I_j$  (column  $j - 1$ ) is visited

```
> meanAbsorptionTime(mcSIS)
  I_1  I_2  I_3  I_4  I_5
24.30 33.45 37.55 39.65 40.85
> absorptionProbabilities(mcSIS)
I_0
I_1  1
I_2  1
I_3  1
I_4  1
I_5  1
```

Sojourn times in compartments

Discrete-time Markov chains

Continuous time Markov chains

Structuration in age

## Continuous-time Markov chains

CTMC similar to DTMC except in way they handle time between events (transitions)

DTMC: transitions occur each  $\Delta t$

CTMC:  $\Delta t \rightarrow 0$  and transition times follow an exponential distribution parametrised by the state of the system

CTMC are roughly equivalent to ODE

## Continuous time Markov chains

ODE  $\leftrightarrow$  CTMC

Simulating CTMC (in theory)

Simulating CTMC (in practice)

Parallelising your code in R

# Converting your compartmental ODE model to CTMC

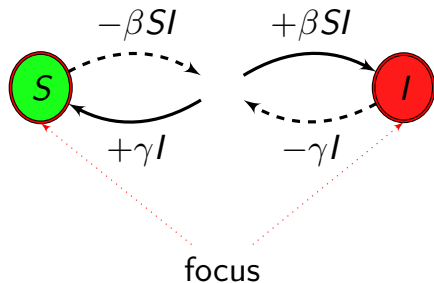
Easy as  $\pi$  :)

- ▶ Compartmental ODE model focuses on flows into and out of compartments
- ▶ ODE model has as many equations as there are compartments
- ▶ Compartmental CTMC model focuses on transitions
- ▶ CTMC model has as many transitions as there are arrows between (or into or out of) compartments

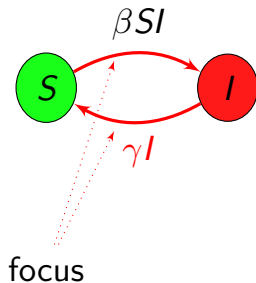


# ODE to CTMC : focus on different components

ODE



CTMC



## SIS without demography

Transition	Effect	Weight	Probability
$S \rightarrow S - 1, I \rightarrow I + 1$	new infection	$\beta SI$	$\frac{\beta SI}{\beta SI + \gamma I}$
$S \rightarrow S + 1, I \rightarrow I - 1$	recovery of an infectious	$\gamma I$	$\frac{\gamma I}{\beta SI + \gamma I}$

States are  $S, I$

## SIS with demography

Transition	Effect	Weight	Probability
$S \rightarrow S + 1$	birth of a susceptible	$b$	$\frac{b}{b+d(S+I)+\beta SI+\gamma I}$
$S \rightarrow S - 1$	death of a susceptible	$dS$	$\frac{dS}{b+d(S+I)+\beta SI+\gamma I}$
$S \rightarrow S - 1, I \rightarrow I + 1$	new infection	$\beta SI$	$\frac{\beta SI}{b+d(S+I)+\beta SI+\gamma I}$
$I \rightarrow I - 1$	death of an infectious	$dI$	$\frac{dI}{b+d(S+I)+\beta SI+\gamma I}$
$S \rightarrow S + 1, I \rightarrow I - 1$	recovery of an infectious	$\gamma I$	$\frac{\gamma I}{b+d(S+I)+\beta SI+\gamma I}$

States are  $S, I$

## Kermack & McKendrick model

Transition	Effect	Weight	Probability
$S \rightarrow S - 1, I \rightarrow I + 1$	new infection	$\beta SI$	$\frac{\beta SI}{\beta SI + \gamma I}$
$I \rightarrow I - 1, R \rightarrow R + 1$	recovery of an infectious	$\gamma I$	$\frac{\gamma I}{\beta SI + \gamma I}$

States are  $S, I, R$

## Continuous time Markov chains

ODE  $\leftrightarrow$  CTMC

Simulating CTMC (in theory)

Simulating CTMC (in practice)

Parallellising your code in R

## Gillespie's algorithm

- ▶ A.k.a. the stochastic simulation algorithm (SSA)
- ▶ Derived in 1976 by Daniel Gillespie
- ▶ Generates possible solutions for CTMC
- ▶ Extremely simple, so worth learning how to implement; there are however packages that you can use (see later)

## Gillespie's algorithm

Suppose system has state  $\mathbf{x}(t)$  with initial condition  $\mathbf{x}(t_0) = \mathbf{x}_0$  and *propensity functions*  $a_i$  of elementary reactions

set  $t \leftarrow t_0$  and  $\mathbf{x}(t) \leftarrow \mathbf{x}_0$

while  $t \leq t_f$

-  $\xi_t \leftarrow \sum_j a_j(\mathbf{x}(t))$

- Draw  $\tau_t$  from  $T \sim \mathcal{E}(\xi_t)$

- Draw  $\zeta_t$  from  $\mathcal{U}([0, 1])$

- Find  $r$ , smallest integer s.t.  $\sum_{k=1}^j a_k(\mathbf{x}(t)) > \zeta_t \sum_j a_j(\mathbf{x}(t)) = \zeta_t \xi_t$

- Effect the next reaction (the one indexed  $r$ )

-  $t \leftarrow t + \tau_t$

## Drawing at random from an exponential distribution

If you do not have an exponential distribution random number generator.. We want  $\tau_t$  from  $T \sim \mathcal{E}(\xi_t)$ , i.e.,  $T$  has probability density function

$$f(x, \xi_t) = \xi_t e^{-\xi_t x} \mathbf{1}_{x \geq 0}$$

Use cumulative distribution function  $F(x, \xi_t) = \int_{-\infty}^x f(s, \xi_t) ds$

$$F(x, \xi_t) = (1 - e^{-\xi_t x}) \mathbf{1}_{x \geq 0}$$

which has values in  $[0, 1]$ . So draw  $\zeta$  from  $\mathcal{U}([0, 1])$  and solve  $F(x, \xi_t) = \zeta$  for  $x$

$$F(x, \xi_t) = \zeta \Leftrightarrow 1 - e^{-\xi_t x} = \zeta$$

$$\Leftrightarrow e^{-\xi_t x} = 1 - \zeta$$

$$\Leftrightarrow \xi_t x = -\ln(1 - \zeta)$$

$$\Leftrightarrow x = \frac{-\ln(1 - \zeta)}{\xi_t}$$



## Gillespie's algorithm (SIS model with only I eq.)

set  $t \leftarrow t_0$  and  $I(t) \leftarrow I(t_0)$

while  $t \leq t_f$

-  $\xi_t \leftarrow \beta(P^* - i)i + \gamma i$

- Draw  $\tau_t$  from  $T \sim \mathcal{E}(\xi_t)$

-  $v \leftarrow [\beta(P^* - i)i, \xi_t] / \xi_t$

- Draw  $\zeta_t$  from  $\mathcal{U}([0, 1])$

- Find  $pos$  such that  $v_{pos-1} \leq \zeta_t \leq v_{pos}$

- switch  $pos$

- 1: New infection,  $I(t + \tau_t) = I(t) + 1$

- 2: End of infectious period,  $I(t + \tau_t) = I(t) - 1$

-  $t \leftarrow t + \tau_t$

## Sometimes Gillespie goes bad

- ▶ Recall that the inter-event time is exponentially distributed
- ▶ Critical step of the Gillespie algorithm:
  - ▶  $\xi_t \leftarrow$  weight of all possible events (*propensity*)
  - ▶ Draw  $\tau_t$  from  $T \sim \mathcal{E}(\xi_t)$
- ▶ So the inter-event time  $\tau_t \rightarrow 0$  if  $\xi_t$  becomes very large for some  $t$
- ▶ This can cause the simulation to grind to a halt

## Example: a birth and death process

- ▶ Individuals born at *per capita* rate  $b$
- ▶ Individuals die at *per capita* rate  $d$
- ▶ Let's implement this using classic Gillespie

(See `simulate_birth_death_CTMC.R` on course GitHub repo)

## Gillespie's algorithm (birth-death model)

- set  $t \leftarrow t_0$  and  $N(t) \leftarrow N(t_0)$   
while  $t \leq t_f$
- $\xi_t \leftarrow (b + d)N(t)$
  - Draw  $\tau_t$  from  $T \sim \mathcal{E}(\xi_t)$
  - $v \leftarrow [bN(t), \xi_t] / \xi_t$
  - Draw  $\zeta_t$  from  $\mathcal{U}([0, 1])$
  - Find  $pos$  such that  $v_{pos-1} \leq \zeta_t \leq v_{pos}$
  - switch  $pos$ 
    - 1: Birth,  $N(t + \tau_t) = N(t) + 1$
    - 2: Death,  $N(t + \tau_t) = N(t) - 1$
  - $t \leftarrow t + \tau_t$

```
b = 0.01 # Birth rate
d = 0.01 # Death rate
t_0 = 0 # Initial time
N_0 = 100 # Initial population

# Vectors to store time and state. Initialise with initial
# condition.
t = t_0
N = N_0

t_f = 1000 # Final time

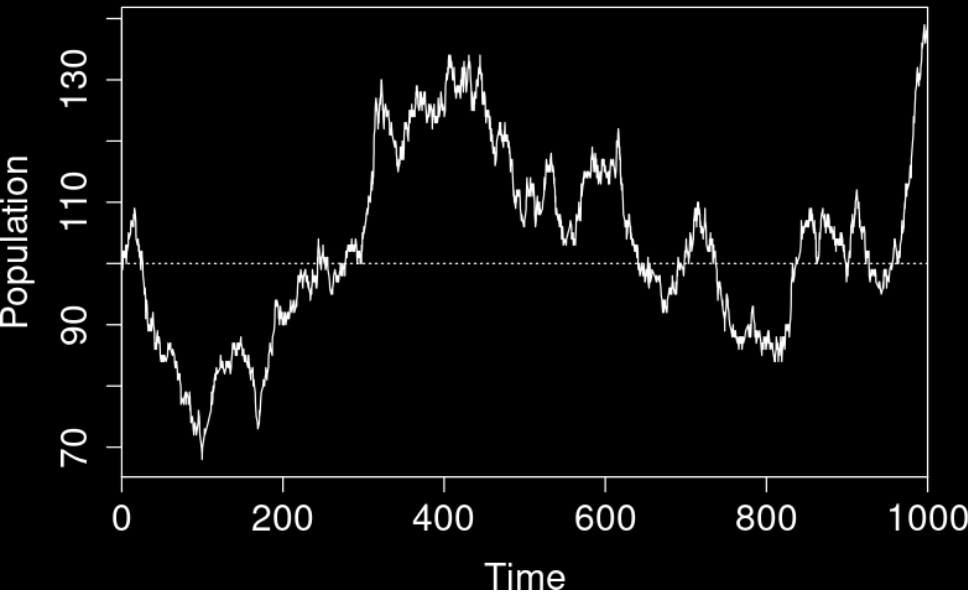
# We'll track the current time and state (could also just check
# last entry in t
# and N, but will take more operations)
t_curr = t_0
N_curr = N_0
```

```

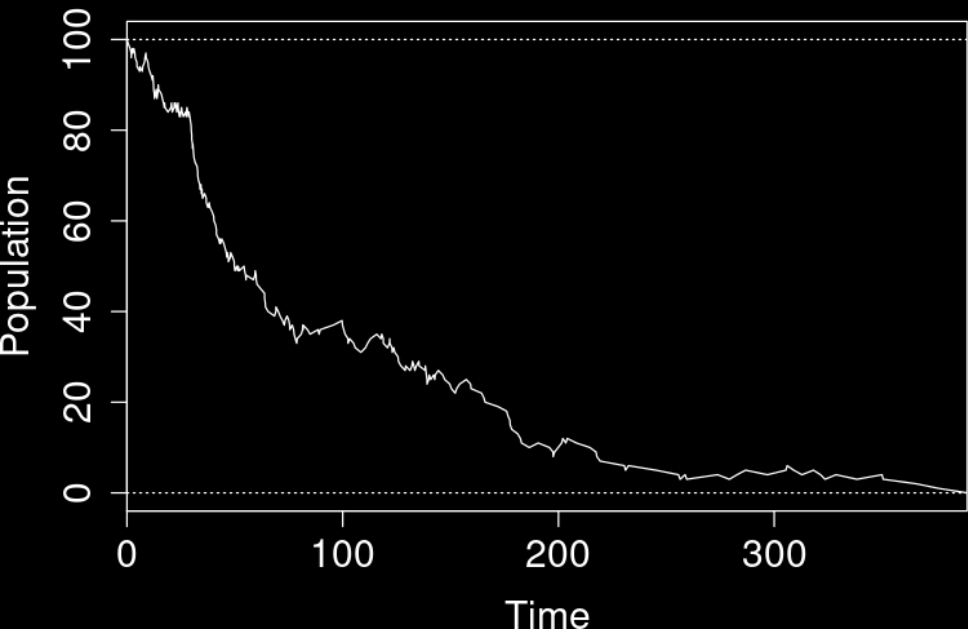
while (t_curr<=t_f) {
  xi_t = (b+d)*N_curr
  # The exponential number generator does not like a rate of 0
  # (when the
  # population crashes), so we check if we need to quit
  if (N_curr == 0) {
    break
  }
  tau_t = rexp(1, rate = xi_t)
  t_curr = t_curr+tau_t
  v = c(b*N_curr, xi_t)/xi_t
  zeta_t = runif(n = 1)
  pos = findInterval(zeta_t, v)+1
  switch(pos,
    {
      N_curr = N_curr+1 # Birth
    },
    {
      N_curr = N_curr-1 # Death
    })
  N = c(N, N_curr)
  t = c(t, t_curr)
}

```

$b=0.01, d=0.01$

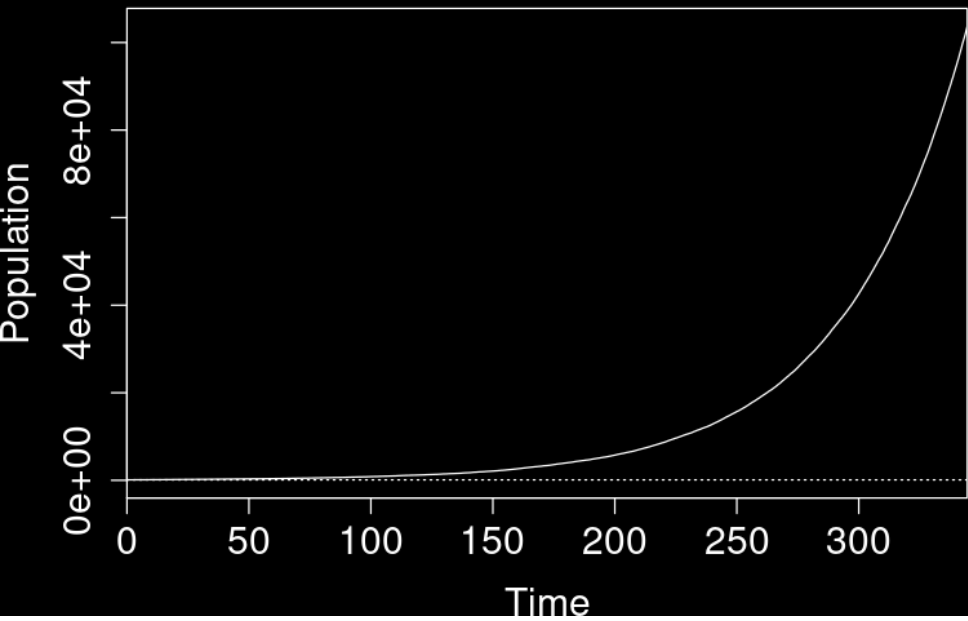


$b=0.01, d=0.02$





$b=0.03, d=0.01$

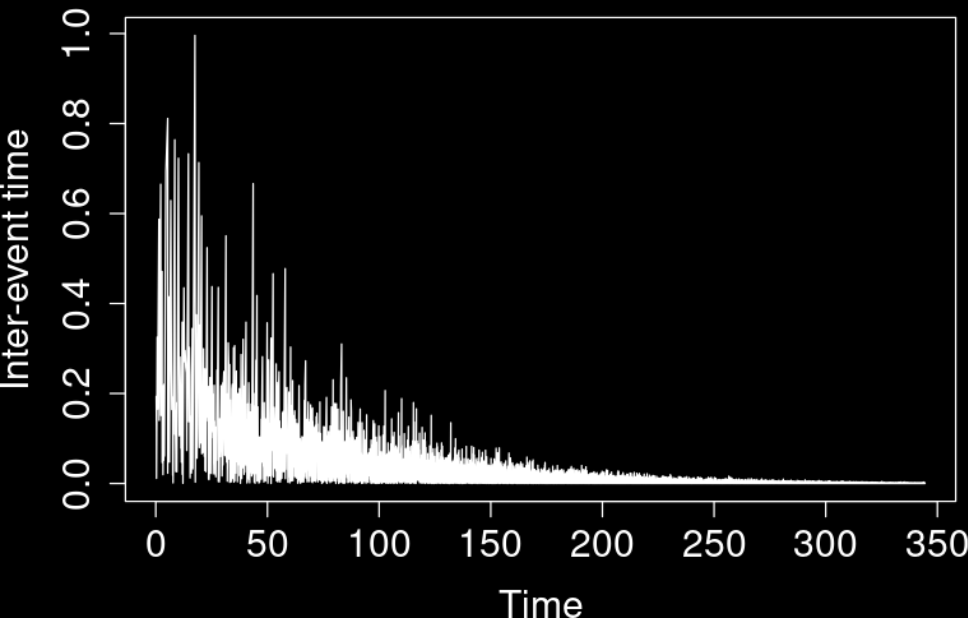


## Last one did not go well

- ▶ Wanted 1000 time units (days?)
- ▶ Interrupted at  $t = 344.4432$  because I lost patience (Penultimate slide: sim stopped because the population went extinct, I did not stop it!)
- ▶ At stop time
  - ▶  $N = 103,646$
  - ▶  $|N| = 208,217$  (and  $|t|$  as well, of course!)
  - ▶ time was moving slowly

```
> tail(diff(t))  
[1] 1.282040e-05 5.386999e-04 5.468540e-04 1.779985e-04 6.737294e-  
-05 2.618084e-04
```

$b=0.03, d=0.01$



## Continuous time Markov chains

ODE  $\leftrightarrow$  CTMC

Simulating CTMC (in theory)

Simulating CTMC (in practice)

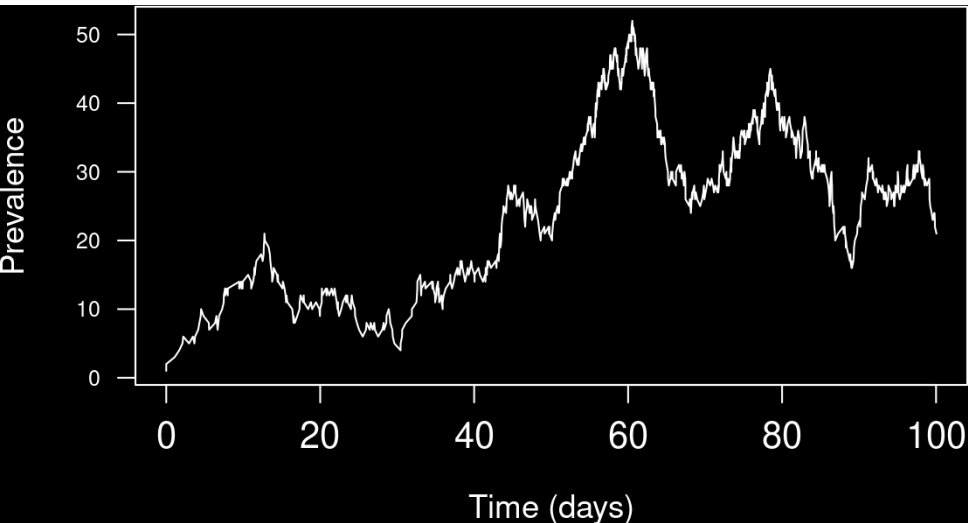
Parallellising your code in R

## Tau-leaping (and packages) to the rescue!

- ▶ *Approximation* method (compared to classic Gillespie, which is exact)
- ▶ Roughly: consider "groups" of events instead of individual events
- ▶ Good news: GillespieSSA2 and adaptivetau, two standard packages for SSA in R, implement tau leaping

## Simulating a CTMC

```
library(GillespieSSA2)
IC <- c(S = (Pop-I_0), I = I_0)
params <- c(gamma = gamma, beta = beta)
reactions <- list(
  reaction("beta*S*I", c(S=-1,I=+1), "new_infection"),
  reaction("gamma*I", c(S=+1,I=-1), "recovery")
)
set.seed(NULL)
sol <- ssa(
  initial_state = IC,
  reactions = reactions,
  params = params,
  method = ssa_exact(),
  final_time = t_f,
)
plot(sol$time, sol$state[, "I"], type = "l",
      xlab = "Time (days)", ylab = "Number infectious")
```



## Continuous time Markov chains

ODE  $\leftrightarrow$  CTMC

Simulating CTMC (in theory)

Simulating CTMC (in practice)

Parallelising your code in R



# Parallelisation

To see multiple realisations: good idea to parallelise, then interpolate results. Write a function, e.g., `run_one_sim` that .. runs one simulation

On the GitHub repo for the course, see

- ▶ `SIS-CTMC-parallel.R`
- ▶ `SIS-CTMC-parallel-multiple-R0.R`

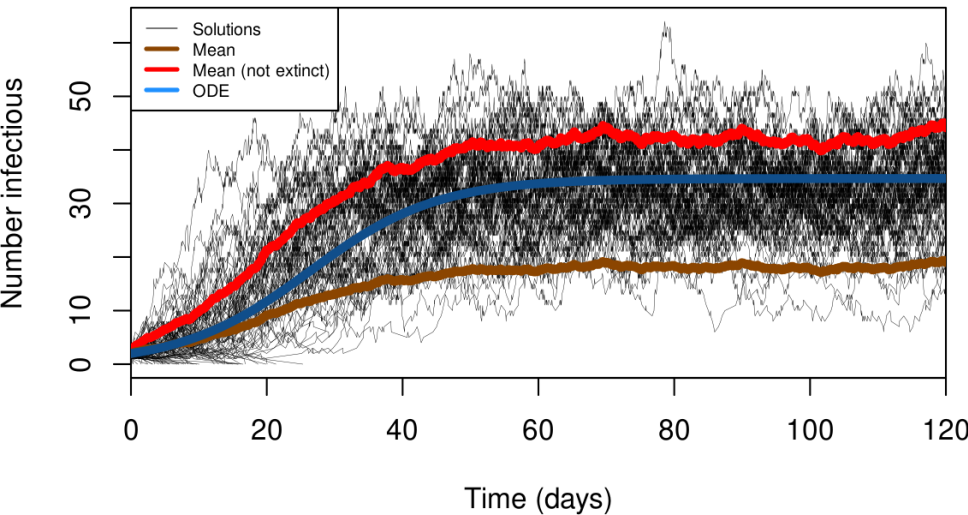
```

run_one_sim = function(params) {
  IC <- c(S = (params$Pop-params$I_0), I = params$I_0)
  params_local <- c(gamma = params$gamma, beta = params$beta)
  reactions <- list(
    # propensity function effects name for reaction
    reaction("beta*S*I", c(S=-1,I=+1), "new_infection"),
    reaction("gamma*I", c(S=+1,I=-1), "recovery")
  )
  set.seed(NULL)
  sol <- ssa(
    initial_state = IC,
    reactions = reactions,
    params = params_local,
    method = ssa_exact(),
    final_time = params$t_f,
    log_firings = TRUE # This way we keep track of events
  )
}

```

```
# Interpolate result (just I will do)
wanted_t = seq(from = 0, to = params$t_f, by = 0.01)
sol$interp_I = approx(x = sol$time, y = sol$state[, "I"], xout
  = wanted_t)
names(sol$interp_I) = c("time", "I")
# Return result
return(sol)
}
```

```
nb_cores <- detectCores()
if (nb_cores > 124) {
  nb_cores = 124
}
cl <- makeCluster(nb_cores)
clusterEvalQ(cl,{
  library(GillespieSSA2)
})
clusterExport(cl,
              c("params",
                "run_one_sim"),
              envir = .GlobalEnv)
SIMS = parLapply(cl = cl,
                X = 1:params$number_sims,
                fun = function(x) run_one_sim(params))
stopCluster(cl)
```



## Benefit of parallelisation

Run the parallel code for 100 sims between 'tictoc::tic()' and 'tictoc::toc()', giving '66.958 sec elapsed', then the sequential version

```
tictoc::tic()
SIMS = lapply(X = 1:params$number_sims,
              FUN = function(x) run_one_sim(params))
tictoc::toc()
```

which gives '318.141 sec elapsed' on a 6C/12T Intel(R) Core(TM) i9-8950HK CPU @ 2.90GHz (4.75× faster) or '12.067 sec elapsed' versus '258.985 sec elapsed' on a 32C/64T AMD Ryzen Threadripper 3970X 32-Core Processor (21.46× faster !)

Sojourn times in compartments

Discrete-time Markov chains

Continuous time Markov chains

Structuration in age

## Age structure

Taking into account age can be important in some cases

- ▶ Demographic characteristics vary with age
- ▶ Interactions are in general more frequent between people of a similar age. They are also more frequent in younger individuals
- ▶ Some diseases attack preferentially younger individuals
- ▶ The immunity of individuals changes with age, so for instance, older people may be more susceptible to some diseases than younger people

This is based on courses given by Jia Li during a Banff summer school in 2004



Lecture Notes in Mathematics

Fred Brauer  
Pauline van den Driessche  
Jianhong Wu (Eds.)

# Mathematical Epidemiology

1945

*Mathematical Biosciences Subseries*



 Springer

## Note on age

**Chronological age**, as a structuring variable, is “easier” than other structuring variables

Indeed, if  $a$  is (chronological) age, then

$$\frac{d}{dt}a = 1$$

## Formulation of an SIR model

Let  $a$  be the age. Assume that natural death and recovery occur at the rates  $\mu$  and  $\gamma$ , respectively, both dependent on  $a$

When an individual is sick, they are subject to disease-induced death at the rate  $\delta(a)$

Governing equations are

$$(\partial_t + \partial_a)S(t, a) = \Lambda(a) - (\mu(a) + \lambda(t, a))S(t, a) \quad (28a)$$

$$(\partial_t + \partial_a)I(t, a) = -(\mu(a) + \gamma(a) + \delta(a))I(t, a) + \lambda(t, a)S(t, a) \quad (28b)$$

$$(\partial_t + \partial_a)R(t, a) = \gamma(a)I(t, a) \quad (28c)$$

Boundary conditions are

$$S(t, a_0) = B \quad (28d)$$

$$I(t, a_0) = 0 \quad (28e)$$

$$R(t, a_0) = 0 \quad (28f)$$

while initial conditions take the form

$$S(0, a) = \Phi(a) \quad (28g)$$

$$I(0, a) = \Psi(a) \quad (28h)$$

$$R(0, a) = 0 \quad (28i)$$

## Force of infection

Transmission  $\lambda(t, a)$  of the disease takes the form

$$\lambda(t, a) = r(a) \int_{a_0}^{\infty} \beta(a, s) \rho(a, s) \frac{I(t, s)}{N(t, s)} ds$$

where

- ▶  $r(a)$  is the number of contacts by individuals of age  $a$  per unit time
- ▶  $\beta(a, s)$  is the probability of disease transmission to a susceptible of age  $a$  by an infectious of age  $s$
- ▶  $\rho(a, s)$  is the meeting rate between people of age  $a$  and people of age  $s$
- ▶  $N(t, a) = S(t, a) + I(t, a) + R(t, a)$  is the distribution of total population

To simplify, assume that  $\beta(a, s)$  is separable

$$\beta(a, s) = f(a)g(s)$$

where  $f(a)$  is the susceptibility of individuals aged  $a$  and  $g(s)$  is the force of infection of individuals aged  $s$

Then

$$\lambda(t, a) = r(a)f(a) \int_{a_0}^{\infty} g(s)\rho(a, s) \frac{I(t, s)}{N(t, s)} ds \quad (29)$$

## Analysis of the SIR model

We seek the DFE by setting  $I = 0$

We find  $(S, I, R) = (S^0(a), 0, 0)$  with

$$S^0(a) = Be^{-M(a)} + e^{-M(a)} \int_{a_0}^a e^{M(x)} \Lambda(x) dx$$

where

$$M(a) = \int_{a_0}^a \mu(s) ds$$

Consider the perturbed solution  $u(t, a) = S(t, a) - S^0(a)$ . Assume that the meeting rate  $\rho$  is also separable,

$$\rho(a, s) = p_1(a)p_2(s)$$

Then

$$\tilde{\lambda}(t, a) := r(a)f(a)p_1(a) \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} I(t, s) ds \simeq \lambda(t, a)$$

and we obtain the linearisation

$$(\partial_t + \partial_a)u = -\mu(a)u - \tilde{\lambda}(t, a)S^0(a)$$

$$(\partial_t + \partial_a)I = -(\mu(a) + \gamma(a) + \delta(a))I + \tilde{\lambda}(t, a)S^0(a)$$

$$(\partial_t + \partial_a)R = \gamma(a)I$$



Let

$$u(t, a) = \tilde{u}(a)e^{c(t-a)} \quad l(t, a) = \tilde{l}(a)e^{c(t-a)}$$

and denote

$$b(a) = S^0(a)r(a)f(a)p_1(a) \quad W = \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} e^{-cs} \tilde{l}(s) ds$$

Then

$$\frac{d\tilde{u}(a)}{da} = -\mu(a)\tilde{u}(a) - b(a)e^{ca}W$$

$$\frac{d\tilde{l}(a)}{da} = -(\mu(a) + \gamma(a))\tilde{l}(a) + b(a)e^{ca}W$$

$$\tilde{l}(a) = We^{-M(a)-\Gamma(a)} \int_{a_0}^{\infty} e^{M(s)+\Gamma(s)} b(s)e^{cs} ds$$

where  $\Gamma(a) = \int_{a_0}^a \gamma(s) ds$

Therefore

$$W = W \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} e^{-M(s)-\Gamma(s)} \int_{a_0}^s e^{M(v)+\Gamma(v)} b(v)e^{-c(s-v)} dv ds$$

Let then

$$H(c) := \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} e^{-M(s)-\Gamma(s)} \int_{a_0}^s e^{M(v)+\Gamma(v)} b(v) e^{-c(s-v)} dv ds$$

We seek roots of the characteristic equation  $H(c) = 1$

We have

$$\frac{dH(c)}{dc} = - \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} e^{-M(s)-\Gamma(s)} \int_{a_0}^s (s-v) e^{M(v)+\Gamma(v)} b(v) e^{-c(s-v)} dv ds < 0$$

implying that  $H(c)$  is a decreasing function

► Let  $c^*$  be a real solution to  $H(c) = 1$ . If  $H(0) > 1$ , then  $c > 0$ , whereas if  $H(0) < 1$ ,  $c < 0$

► Suppose that  $c^* = \alpha + i\beta$  is a complex root of  $H(c) = 1$ . Then

$$\operatorname{Re} H(c) = \int_{a_0}^{\infty} \frac{g(s)p_2(s)}{S^0(s)} e^{-M(s)-\Gamma(s)} \int_{a_0}^s e^{M(v)+\Gamma(v)} b(v) e^{-\alpha(s-v)} \cos \beta(s-v) dv ds$$

As a consequence,  $H(0) < 1 \implies \alpha < 0$

So  $H(0) = 1$  is a threshold and we take  $\mathcal{R}_0 = H(0)$

## Analysis using semigroups: SIA model

To illustrate the use of the semigroup method in this context, we consider an SIA model describing the evolution of HIV/AIDS

The model is almost equivalent to (28), with a few differences

The  $I$  compartment contains individuals bearing HIV, but not yet in the AIDS stage

The rate  $\gamma(a)$  represents the progression towards the AIDS stage

The AIDS stage is represented by compartment  $A$ , where individuals are subject to a specific mortality rate

$$(\partial_t + \partial_a)S(t, a) = \Lambda(a) - (d(a) + \lambda(t, a))S(t, a) \quad (30a)$$

$$(\partial_t + \partial_a)I(t, a) = -(d(a) + \gamma(a))I(t, a) + \lambda(t, a)S(t, a) \quad (30b)$$

$$(\partial_t + \partial_a)A(t, a) = \gamma(a)A(t, a) - (d(a) + \delta(a))A(t, a) \quad (30c)$$

Assume

$$\lambda(t, a) = h(a) \int_{a_0}^{\infty} \rho(a, a') \frac{I(t, a')}{T(t, a')} da' \quad (30d)$$

where  $T(t, a') = S(t, a') + I(t, a')$

An individual in AIDS stage no longer has contacts. Therefore the dynamics of  $S$  and  $I$  do not depend on the dynamics of  $A$ , and we consider the system consisting of the first two variables

Let  $\omega$  be the maximum age. The system in proportions takes the form

$$x := \frac{S}{T} \quad y := \frac{I}{T}$$

As we are only considering  $S$  and  $I$ , we have  $x + y = 1$  and the system reads

$$(\partial_t + \partial_a)y(t, a) = (1 - y)(-\gamma(a)y + \lambda(t, a)) \quad (31a)$$

$$\lambda(t, a) = h(a) \int_0^\omega p(a, a')y(t, a')da' \quad (31b)$$

Let  $X = \{f \in L^1(0, \omega)\}$ . Define

$$(Af)(a) := -\frac{d}{da}f(a), \quad f \in D(A)$$

with  $D(A) = \{f \in X, f \text{ is absolutely continuous, } f(0) = 0\}$ , and

$$F(f)(a) \equiv (1 - f(a)) \left( -\gamma(a)f(a) + h(a) \int_0^\omega p(a, a')f(a')da' \right)$$

an operator from  $X \rightarrow X$

Let  $\Omega = \{f \in X, 0 \leq f \leq 1 \text{ a.e.}\}$ . Then (31) takes the form

$$\begin{aligned} \frac{dy}{dt} &= Ay + F(y) \\ y(0) &= y_0 \in \Omega \end{aligned}$$



Let

$$(\mathcal{B}f)(a) = -\frac{df(a)}{da} - \gamma(a)f(a) \quad (\mathcal{P}f)(a) = h(a) \int_0^\omega \rho(a, a')f(a')da'$$

We have

$$(\partial_t + \partial_a)y = -\gamma(a)y + h(a) \int_0^\omega \rho(a, a')y(t, a')da' \Leftrightarrow \frac{dy}{dt} = (\mathcal{B} + \mathcal{P})y$$

$\mathcal{B} + \mathcal{P}$  generates a  $C_0$ -semigroup  $T(t)$ ,  $t \geq 0$ , which is eventually uniformly continuous

The resolvent of  $\mathcal{B} + \mathcal{P}$  is

$$R(\lambda; \mathcal{B} + \mathcal{P}) = (S_\lambda - I)^{-1} G$$

with

$$(Gf)(a) = \int_0^a e^{-\lambda(a-\sigma)} \frac{\Gamma(a)}{\Gamma(\sigma)} f(\sigma) d\sigma$$

$$(S_\lambda f)(a) = \int_0^\omega \int_0^a e^{-\lambda(a-\sigma)} \frac{\Gamma(a)}{\Gamma(\sigma)} \rho(\sigma, \xi) d\sigma f(\xi) d\xi$$

where we denoted

$$\Gamma(a) = \exp \left( - \int_0^a \gamma(a') da' \right)$$

$\mathcal{R}_0$ 

$\mathcal{R}_0$  is the spectral radius of the operator

$$(Sf)(a) = \int_0^\omega \int_0^a \frac{\Gamma(a)}{\Gamma(\sigma)} h(\sigma) p(\sigma, \xi) d\sigma f(\xi) d\xi$$

## Pair formation

$\rho(t, a, a')$  proportion of partners of an individual aged  $a$  who are aged  $a'$

$r(t, a)$  mean number of partners of an individual aged  $a$

$T(t, a)$  total number of individuals aged  $a$

The following conditions must hold

- ▶  $0 \leq \rho \leq 1$
- ▶  $\int_0^\infty \rho(t, a, a') da' = 1$
- ▶  $\rho(t, a, a') r(t, a) T(t, a) = \rho(t, a', a) r(t, a') T(t, a')$
- ▶  $r(t, a) T(t, a) r(t, a') T(t, a') = 0 \Rightarrow \rho(t, a, a') = 0$